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# ON THE ASYMPTOTIC NORMALITY FOR NONPARAMETRIC SEQUENTIAL DENSITY ESTIMATION\*

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#### Abstract

Let  $f_n(x)$  be a recursive kernel estimator of a probability density function f at a point x. We show that if N(t) is a sequence of positive integer-valued random variables and  $\pi(t)$  a sequence of positive numbers with  $N(t)/\pi(t) \rightarrow \theta$  in probability as  $t \rightarrow \infty$ , where  $\theta$  is a positive discrete random variable, then  $(N(t)h_{N(t)}^p)^{1/2}(f_{N(t)}(x)-f(x))$  is asymptotically normally distributed under certain conditions.

#### 1. Introduction

Let X be a p-dimensional random vector on a probability space  $(\Omega, \mathcal{B}, P)$  having a probability density function (p, d, f) with respect to the Lebesgue measure on  $R^p$ . There is a vast literature on the problem of estimating the p, d, f (see Devroye and Györfi [3], and Prakasa Rao [9] for example). In particular, estimators have been proposed in some recursive manners by several authors on behalf of the following two advantages: data need not be stored, and the estimators are easily updated when new data become available. In this paper we consider the recursive kernel estimator proposed by the author  $\lceil 4 \rceil$ .

On the other hand, in many practical situations the number of observations N(t) which we observe in a time-interval (0,t] is random. The problem of sequential estimation of the p.d.f. by using positive integer-valued random variables (i.e., stopping rules) were studied by Davies and Wegman [2], Carroll [1], Wegman and Davies [11] and the author [5], for example. Carroll [1] and the author [5] investigated the asymptotic normality of estimates of the p.d.f. under random sample sizes. In this paper we shall show that the asymptotic normality holds for a more general class of positive integer-valued random variables N(t) than the classes of Carroll [1] and the author [5]. We note that the extension to this general class was motivated by the discussion in Rényi [10]. Throughout this paper we consider the estimator  $f_{N(t)}(x)$  of the p.d.f. f(x) based on  $X_1, X_2, \dots, X_{N(t)}$ , which is defined by

$$f_{N(t)}(x) = \sum_{j=1}^{N(t)} a_j \beta_{jN(t)} K_j(x, X_j) + \beta_{0N(t)} K(x), \qquad (1.1)$$

where  $X_1, X_2, \cdots$  are independent observations of  $X_1$ 

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$$K_n(x, y) = h_n^{-p} K((x-y)/h_n)$$
 for  $x, y \in \mathbb{R}^p$ , (1.2)

K is a bounded, integrable, real-valued Borel measurable function on  $R^p$  and  $\{h_n\}$  with  $h_0=h_1$  is a nonincreasing sequence of positive numbers converging to zero,

$$a_n = a/n$$
 for any fixed  $a \in (0, 1]$ , (1.3)

and

$$\beta_{mn} = \begin{cases} \prod_{j=m+1}^{n} (1-a_j) & \text{if } n > m \ge 0\\ 1 & \text{if } n = m \ge 0. \end{cases}$$
 (1.4)

The aim of this paper is to show that under certain conditions  $(N(t)h_{N(t)}^p)^{1/2}(f_{N(t)}(x)-f(x))$  is asymptotically normally distributed. In Section 2 we shall make some preparations and auxiliary results. In Section 3 we shall give our main theorem.

#### 2. Auxiliary Results

In this section we shall make some preparations and auxiliary results. Set

$$\gamma_1=1$$
 and  $\gamma_n=\sum_{j=2}^n(1-a_j)$  for  $n\geq 2$ .

where  $a_n$  is as defined in (1.3). Clearly,

$$\beta_{mn} = \gamma_n \gamma_m^{-1}$$
 for  $n \ge m \ge 1$ . (2.1)

It is known in [4] that

$$L_1 n^{-a} \le \gamma_n \le L_2 n^{-a}$$
 for some constants  $L_1$ ,  $L_2 > 0$  and all  $n \ge 1$  (2.2)

and

$$\beta_{mn} \sim m^a n^{-a}$$
 as  $n \ge m \to \infty$ , (2.3)

where " $\sim$ " means the asymptotic equivalence. For a real-valued functing g let C(g) be the set of continuity points of g. Throughout this paper we assume the function K in Section 1 to satisfy

$$\begin{split} & \int_{R^p} K(u) du = 1 \,, \qquad \int_{R^p} \|u\|_p^2 |K(u)| \, du < \infty \\ & \int_{R^p} u_i K(u) du = 0 \qquad \text{for} \quad i = 1, \, \cdots, \, p \quad \text{with} \quad u = (u_1, \, \cdots, \, u_p) \end{split}$$

and

$$||u||_p^p |K(u)| \to 0$$
 as  $||u||_p \to \infty$ ,

where  $\|\cdot\|_p$  denotes the Euclidean norm on  $\mathbb{R}^p$ . On the sequence  $\{h_n\}$  in Section 1 we shall impose some or all of the following conditions: For a fixed  $a \in (0, 1]$ ,

- (H1)  $nh_n^p \uparrow \infty$  as  $n \to \infty$ ,
- (H2)  $n^{1-2a}h_n^p \to 0$  as  $n \to \infty$ ,
- (H3)  $n^{1-2a}h_n^p\sum_{j=1}^n j^{2(a-1)}h_j^{-p} \rightarrow \beta$  as  $n\rightarrow\infty$  for some constant  $\beta>0$ ,

$$({\rm H4}) \qquad n^{3/2-3a} \, h_n^{3p/2} \sum_{i=1}^n j^{3(a-1)} h_j^{-2p} \! \to \! 0 \qquad \text{as} \quad n \! \to \! \infty \, ,$$

(H5) 
$$(n^{1-2a}h_n^p)^{1/2}\sum_{j=1}^n j^{a-1}h_j^2 \to 0$$
 as  $n \to \infty$ ,

(H6) For any  $\varepsilon > 0$  there exists a positive constant  $\delta = \delta(\varepsilon)$  such that  $|n/m-1| < \delta$  implies  $|h_n/h_m-1| < \varepsilon$ .

EXAMPLE.

Let

$$h_n = n^{-r/p}$$
 with  $\max\{p/(p+4), 1-2a\} < r < 1$ .

Then  $\{h_n\}$  satisfies (H1) $\sim$ (H6) with  $\beta=(2a+r-1)^{-1}$ . Throughout this paper  $C, C_1, C_2, \cdots$  denote appropriate positive constants. The following lemma can be found in the author  $\lceil 6 \rceil$ .

LEMMA 2.1. Let  $\{h_n\}$  be a sequence of positive numbers converging to zero. Suppose that k is a bounded, integrable, real-valued Borel measurable function on  $R^p$  satisfying

$$||u||_p^p |k(u)| \rightarrow 0$$
 as  $||u||_p \rightarrow \infty$ .

Let g be an integrable, real-valued Borel measurable function on  $R^p$ . Then for each point  $x \in C(g)$ ,

$$\int_{\mathbb{R}^p} h_n^{-p} \, k((x-u)/h_n) g(u) d\, u \to g(x) \int_{\mathbb{R}^p} k(u) d\, u \qquad \text{as} \quad n \to \infty$$

and

$$\sup_{n\geq 1} \int_{\mathbb{R}^p} h_n^{-p} |k((x-u)/h_n)| |g(u)| du \leq C,$$

where C may depend on x.

LEMMA 2.2. Let a constant  $a \in (0, 1]$  be given. Suppose that a sequence of positive numbers  $\{h_n\}$  converging to zero satisfies (H2), (H3) and (H4). Let  $\{Z_n\}$  be a sequence of independent random variables with  $EZ_n=0$ . Assume that

 $h_n^p E Z_n^2 \to \xi$  as  $n \to \infty$  for some constant  $\xi > 0$ 

and

$$h_n^{2p}E|Z_n|^3 \leq C$$
 for all  $n \geq 1$ .

Then,

$$(nh_n^p)^{1/2} \sum_{j=1}^n a_j \beta_{jn} Z_j \xrightarrow{L} N(0, B)$$
 as  $n \to \infty$  (in law),

where  $B=a^2\beta\xi$  (>0), and  $a_n$  and  $\beta_{mn}$  are as defined in (1.3) and (1.4), respectively. PROOF. It was shown in Lemma 2.2 of [6] that

$$\sum_{j=1}^n (j^2 \gamma_j^2 h_j^p)^{-1} \sim \beta (n h_n^p \gamma_n^2)^{-1}$$
 as  $n \to \infty$ ,

which, together with (2.1), (2.2), (H2), the assumption of  $EZ_n^2$  and the Toeplitz lemma (see Loève [7], page 238), implies that

$$nh_n^p \sum_{j=1}^n j^{-2} \beta_{jn}^2 E Z_j^2 \to \beta \xi$$
 as  $n \to \infty$ . (2.4)

Set

$$U_n = a_n \gamma_n^{-1} Z_n$$
,  $S_n = \sum_{j=1}^n U_j$  and  $s_n^2 = \text{Var}(S_n) = a^2 \sum_{j=1}^n j^{-2} \gamma_j^{-2} E Z_j^2$ .

From (2.4)

$$s_n^2 \sim B(nh_n^p \gamma_n^2)^{-1}$$
 as  $n \to \infty$ . (2.5)

By the assumption of  $E|Z_n|^3$  and (2.2) we get that

$$E|U_n|^3 \leq C_1 n^{3(a-1)} h_n^{-2p}$$
 for all  $n \geq 1$ ,

which, together with (2.2), (2.5) and (H4), yields that

$$s_n^{-3} \sum_{j=1}^n E|U_j|^3 \to 0$$
 as  $n \to \infty$ .

Thus by the Liapounov theorem we have

$$S_n^{-1}S_n \longrightarrow N(0, 1)$$
 as  $n \to \infty$ . (2.6)

From (2.5) and (2.6) we obtain

$$(nh_n^p)^{1/2} \sum_{j=1}^n a_j \beta_{jn} Z_j = (nh_n^p \gamma_n^2 s_n^2)^{1/2} s_n^{-1} S_n \xrightarrow{L} N(0, B)$$
 as  $n \to \infty$ .

This completes the proof.

We shall give a definition of the smoothness of a function g.

DEFINITION. Let g be a real-valued function on  $R^p$ . We say that the function g belongs to the class  $\mathcal{M}_p$  (abbreviated as  $g \in \mathcal{M}_p$ ) if there exist bounded, continuous second partial derivatives  $\partial^2 g(x)/\partial x_i \partial x_j$  on  $R^p$  for all  $i, j=1, \dots, p$ .

LEMMA 2.3. Assume  $g \in \mathcal{M}_p$ . Suppose that k is a real-valued Borel measurable function on  $R^p$  satisfying

$$\int_{\mathbb{R}^p} u_i k(u) du = 0 \quad \text{for } i=1, \dots, p \quad \text{with} \quad u = (u_1, \dots, u_p)$$

$$\int_{\mathbb{R}^p} ||u||_p^2 |k(u)| du < \infty.$$

and

Then there exists a positive constant C not depending on h such that

$$\sup_{x\in R^p}\left|\int_{R^p}k(u)\{g(x-hu)-g(x)\}du\right|\leq Ch^2\quad \text{for all}\quad h>0.$$

The proof of this lemma is omitted because it is easily shown by the Taylor theorem.

The following proposition shows the asymptotic normality of  $(nh_n^p)^{1/2}(f_n(x)-f(x))$ . Proposition 2.4. Let  $\{h_n\}$  satisfy (H2) $\sim$ (H5). Assume  $f \in \mathcal{M}_p$ . Then for each point x with f(x)>0,

$$(nh_n^p)^{1/2}(f_n(x)-f(x)) \longrightarrow N(0, \sigma^2(x))$$
 as  $n \to \infty$ .

where

$$\sigma^2(x) = a^2 \beta f(x) \int_{\mathbb{R}^p} K^2(u) du.$$

PROOF. Let any x with f(x) > 0 be fixed. Set

$$Z_n = K_n(x, X_n) - EK_n(x, X_n)$$
 and  $\delta_n = EK_n(x, X_n) - f(x)$ .

Then, replacing N(t) in (1.1) by n we get

$$(nh_n^p)^{1/2}(f_n(x)-f(x)) = (nh_n^p)^{1/2}\beta_{0n}(K(x)-f(x))+(nh_n^p)^{1/2}\sum_{j=1}^n a_j\beta_{jn}Z_j+(nh_n^p)^{1/2}\sum_{j=1}^n a_j\beta_{jn}\delta_j.$$
 (2.7)

From (2.2) and (H2) the first term in the right hand side of (2.7) converges to zero as n tends to infinity. In view of Lemma 2.3, (2.1), (2.2) and (H5) the last term in the right hand side of (2.7) converges to zero as n tends to infinity. Thus the proposition will be proved if we show that

$$(nh_n^p)^{1/2} \sum_{j=1}^n a_j \beta_{jn} Z_j \longrightarrow N(0, \sigma^2(x))$$
 as  $n \to \infty$ . (2.8)

From Lemma 2.1 and (1.2)

$$h_n^p E Z_n^2 \rightarrow f(x) \int_{\mathbb{R}^p} K^2(u) du \ (>0)$$
 as  $n \rightarrow \infty$ .

By the Hölder inequality and Lemma 2.1 we have that  $h_n^{2p} E |Z_n|^8 \le C_1$  for all  $n \ge 1$ . Since all the conditions of Lemma 2.2 are satisfied, the relation (2.8) holds. This completes the proof.

The next lemma was provided by Rényi [10].

LEMMA 2.5. Let  $\{Y_n\}$  be a sequence of independent random variables defined on a probability space  $(\Omega, A, P)$  such that putting

$$S_n = \frac{1}{B_n} \sum_{j=1}^n Y_j$$
 where  $B_n \to \infty$ 

the random variable  $S_n$  converges in law to a random variable with the distribution function F. Then for any event  $A \in \mathcal{A}$  with P(A) > 0 the conditional probability  $P\{S_n < x \mid A\}$  tends to F(x) for every  $x \in C(F)$ .

#### 3. Main Result

In this section we shall show the asymptotic normality of  $(N(t)h_{N(t)}^p)^{1/2}(f_{N(t)}(x)-f(x))$ . Let [b] denote the largest integer not greater than b. For any fixed  $x \in \mathbb{R}^p$  set

$$U_n^{(1)} = K_n(x, X_n) - EK_n(x, X_n), \qquad U_n^{(2)} = EK_n(x, X_n) - f(x),$$

$$S_n = \sum_{j=1}^n a_j \beta_{jn} \{ K_j(x, X_j) - f(x) \}, \qquad V_n = (nh_n^p)^{1/2} S_n \quad \text{for} \quad n \ge 1,$$
(3.1)

and  $S_0 = V_0 = 0$ . It is clear from (2.7) that

$$(nh_n^p)^{1/2}(f_n(x)-f(x))=V_n+(nh_n^p)^{1/2}\beta_{0n}(K(x)-f(x)) \quad \text{for} \quad n\ge 1.$$
 (3.2)

Now, we shall give the condition on N(t). For any  $t \in (0, \infty)$  let N(t) be a positive integer-valued random variable defined on the probability space  $(\Omega, \mathcal{B}, P)$  given in Section 1.

DEFINITION. A sequence of positive integer-valued random variables N(t) is said to satisfy Condition A if there exist a positive random variable  $\theta$  defined on  $(\mathcal{Q}, \mathcal{B}, P)$  having a discrete distribution and a sequence of positive numbers  $\pi(t)$  with  $\pi(t) \rightarrow \infty$  as  $t \rightarrow \infty$  such that

$$N(t)/\pi(t) \xrightarrow{P} \theta$$
 as  $t \to \infty$  (in probability).

Here, by the positive random variable  $\theta$  having a discrete distribution we mean that there exists a sequence of positive numbers  $l_k$   $(k=1, 2, \cdots)$  (k may be finite or infinite) such that

$$\sum_{k=1}^{\infty} p_k = 1 \quad \text{where} \quad p_k = P\{\theta = l_k\} > 0.$$
 (3.3)

Throughout this section  $\pi(t)$  and  $\theta$  are as given in the above definition.

REMARK. The stopping rules N(t) treated by Carroll [1] and the author [5] satisfy Condition A with  $P\{\theta=1\}=1$ .

LEMMA 3.1. Let  $\{h_n\}$  be a nonincreasing sequence of positive numbers converging to zero and satisfy (H1) and (H5). Let  $\{\delta_n\}$  be a sequence of real numbers satisfying  $|\delta_n| \le C_1 h_n^2$  for all  $n \ge 1$ . Suppose that  $\{Z_n\}$  is a sequence of independent random variables satisfying

$$EZ_n=0$$
,  $h_n^p EZ_n^2 \leq C_2$  and  $nh_n^p \sum_{i=1}^n a_i^2 \beta_{jn}^2 EZ_j^2 \leq C_2$  for all  $n \geq 1$ .

Set

$$W_n = \sum_{j=1}^n a_j \beta_{jn} Z_j + \sum_{j=1,j=1}^n a_j \beta_{jn} \delta_j$$
.

If N(t) satisfies Condition A then

$$(N(t)h_{N(t)}^{p})^{1/2}(W_{N(t)}-W_{\lceil\theta\pi(t)\rceil}) \xrightarrow{p} 0 \quad \text{as} \quad t \to \infty.$$

The proof of this lemma is deferred to Appendix. We shall now state our result.

THEOREM. Assume  $f \in \mathcal{M}_p$ . Let  $\{h_n\}$  satisfy (H1) $\sim$ (H6). Suppose that N(t) satisfies Condition A. Then for each point x with f(x)>0,

$$(N(t)h_{N(t)}^p)^{1/2}(f_{N(t)}(x)-f(x)) \xrightarrow{L} N(0, \sigma^2(x)) \quad \text{as} \quad t {\to} \infty,$$

where

$$\sigma^2(x) = a^2 \beta f(x) \int_{\mathbb{R}^p} K^2(u) du.$$

PROOF. For simplicity put N=N(t). Let any x with f(x)>0 be fixed. First we shall show that

$$V_{[\theta_{\pi(t)}]} \xrightarrow{L} N(0, \sigma^2(x))$$
 as  $t \to \infty$ . (3.4)

Since by (2.2) and (H2)

$$(nh_n^p)^{1/2}\beta_{0n} \to 0 \quad \text{as} \quad n \to \infty, \tag{3.5}$$

it follows from Proposition 2.4 and (3.2) that

$$V_n \xrightarrow{L} N(0, \sigma^2(x))$$
 as  $n \to \infty$ .

Hence by Lemma 2.5 we get that for any fixed k

$$P\{V_n < y \mid \theta = l_k\} \rightarrow F(y)$$
 as  $n \rightarrow \infty$  for each  $y \in R$ , (3.6)

where F denotes the distribution function of  $N(0, \sigma^2(x))$ . Let any  $\epsilon > 0$  be fixed. From

(3.3) there exists a positive integer  $k_0$  such that

$$\sum_{k=k_0+1}^{\infty} p_k < \varepsilon. \tag{3.7}$$

Fix any  $y \in R$ . By (3.3) and (3.7)

$$|P\{V_{[\theta\pi(t)]} < y\} - F(y)| < \sum_{k=1}^{k_0} |P\{V_{n(k,t)} < y | \theta = l_k\} - F(y)| + \varepsilon \text{ for any } t \in (0, \infty), (3.8)$$

where  $n(k, t) = [l_k \pi(t)]$ . Hence, in view of (3.6) and (3.8) we obtain (3.4). From (3.2) it is clear that

$$(Nh_N^p)^{1/2} (f_N(x) - f(x)) = V_N + (Nh_N^p)^{1/2} \beta_{0N}(K(x) - f(x)).$$
 (3.9)

Since Condition A implies that  $N \longrightarrow \infty$  as  $t \to \infty$ , by use of (3.5)

$$(Nh_N^p)^{1/2}\beta_{0N} \xrightarrow{P} 0 \quad \text{as} \quad t \to \infty.$$
 (3.10)

Thus, in view of (3.9) and (3.10), in order to prove the theorem it suffices to show that

$$V_N \xrightarrow{L} N(0, \sigma^2(x))$$
 as  $t \to \infty$ . (3.11)

From (3.1)

$$V_N = V_{\lceil \theta \pi(t) \rceil} + (Nh_N^p)^{1/2} (S_N - S_{\lceil \theta \pi(t) \rceil}) + V_{\lceil \theta \pi(t) \rceil} \{ (Nh_N^p / (\lceil \theta \pi(t) \rceil h_{\lceil \theta \pi(t) \rceil}^p))^{1/2} - 1 \}.$$

Hence, taking account of (3.4), in order to show (3.11) it suffices to prove that

$$(Nh_N^p)^{1/2}(S_N - S_{(\theta\pi(t))}) \xrightarrow{p} 0 \quad \text{as} \quad t \to \infty$$
 (3.12)

and

$$V_{[\theta\pi(t)]}\{(Nh_N^p/([\theta\pi(t)]h_{[\theta\pi(t)]}^p))^{1/2}-1\} \longrightarrow 0 \quad \text{as} \quad t \to \infty.$$
 (3.13)

First we shall show (3.13). Condition A implies that

$$N/[\theta\pi(t)] \xrightarrow{P} 1$$
 as  $t\to\infty$ ,

which, together with (H6), yields that

$$Nh_N^p/([\theta\pi(t)]h_{[\theta\pi(t)]}^p) \xrightarrow{P} 1$$
 as  $t \to \infty$ . (3.14)

Thus, by virtue of (3.4) and (3.14) we obtain (3.13). Finally, we shall show (3.12). From (3.1) we get

$$S_n = \sum_{i=1}^{2} \sum_{j=1}^{n} a_j \beta_{jn} U_j^{(i)}. \tag{3.15}$$

By Lemma 2.3

$$|U_n^{(2)}| \le C_1 h_n^2$$
 for all  $n \ge 1$ . (3.16)

From Lemma 2.1

$$h_n^p E\{(U_n^{(1)})^2\} \le h_n^p E K_n^2(x, X_n) \le C_2 \quad \text{for all} \quad n \ge 1.$$
 (3.17)

(H3) implies that

$$n^{1-2a}h_n^p \sum_{i=1}^n j^{2(a-1)}h_j^{-p} \le C_3$$
 for all  $n \ge 1$ ,

which, together with (2.1), (2.2) and (3.17), yields that

$$nh_n^p \sum_{j=1}^n a_j^2 \beta_{jn}^2 E\{(U_j^{(1)})^2\} \le C_4$$
 for all  $n \ge 1$ . (3.18)

Thus, combining Lemma 3.1 and (3.15) to (3.18) we obtain (3.12). This completes the proof.

#### **Appendix**

PROOF OF LEMMA 3.1. Let any positive numbers  $\varepsilon$  and  $\xi$  be fixed. From (3.3) there exists a positive integer  $k_0$  such that

$$\sum_{k=k_0+1}^{\infty} p_k < \xi/4. \tag{A.1}$$

Fix a positive constant  $C_1$ , which will be chosen later. Choose  $\rho(0 < \rho < 1/2)$  such that

$$C_1 \varepsilon^{-2} \{1 - ((1-\rho)/(1+\rho))^a\}^2 < \xi/(8k_0)$$
 (A.2)

and

$$C_1 \varepsilon^{-2} \rho < \xi/(8k_0). \tag{A.3}$$

For each  $t \in (0, \infty)$  let n(t) be a nonnegative integer with  $n(t) \to \infty$  as  $t \to \infty$ . Set

$$M_1 = [(1-\rho)n(t)]$$
 and  $M_2 = [(1+\rho)n(t)]$ . (A.4)

By virtue of (2.3) it is easy to show that for all  $t \ge$  some  $t_0$ 

$$1 \le M_1, \ 1 \le M_2 - M_1 < 2\rho M_2, \ M_2 / M_1 < 3$$
 (A.5)

and

$$(1-\beta_{M_1M_2})^2 < 2\{1-((1-\rho)/(1+\rho))^a\}^2. \tag{A.6}$$

By the assumption of  $\delta_n$  and (A.5)

$$\begin{split} (M_{2}h_{M_{2}}^{p})^{1/2} \max_{M_{1} \leq i \leq M_{2}} \left| \sum_{j=1}^{i} a_{j} \beta_{ji} \delta_{j} \right| &\leq C_{2} (M_{2}/M_{1})^{a} M_{2}^{1/2-a} h_{M_{2}}^{p/2} \sum_{j=1}^{M_{2}} j^{a-1} h_{j}^{2} \\ &\leq C_{3} M_{2}^{1/2-a} h_{M_{2}}^{p/2} \sum_{j=1}^{M_{2}} j^{a-1} h_{j}^{2} \quad \text{for } t \geq t_{0}. \end{split} \tag{A.7}$$

From (H5) there exists a positive integer  $n_0$  such that

$$n^{1/2-a}h_n^{p/2} \sum_{j=1}^n j^{a-1}h_j^2 < \varepsilon/(8C_3)$$
 for all  $n \ge n_0$ . (A.8)

As  $M_2 \ge n_0$  for all  $t \ge$  some  $t_1 (\ge t_0)$ , (A.7) and (A.8) yield that

$$(M_2 h_{M_2}^p)^{1/2} \max_{M_1 \le i \le M_2} \left| \sum_{j=1}^i a_j \beta_{ji} \delta_j \right| < \varepsilon/8 \quad \text{for all } t \ge t_1.$$
 (A.9)

Set  $n(k, t) = [l_k \pi(t)]$  for  $k=1, 2, \cdots$ . For simplicity put N=N(t). It is clear that

$$P\{(Nh_N^p)^{1/2}|W_N - W_{[\theta_\pi(t)]}| \ge \varepsilon\} \le I_1(t) + I_2(t),$$
 (A.10)

where

$$I_1(t) = \sum_{k=1}^{\infty} P\{(Nh_N^p)^{1/2} | W_N - W_{n(k,t)} | \geq \varepsilon, |N - n(k,t)| < \rho \, n(k,t), \; \theta = l_k \}$$

and

$$I_2(t) = P\{|N - [\theta \pi(t)]| \ge \rho [\theta \pi(t)]\}.$$

Condition A implies that

$$I_2(t) < \xi/2$$
 for all  $t \ge \text{some } t_2$ . (A.11)

From (A.1)

$$I_{1}(t) < \sum_{k=1}^{k_{0}} P\{(Nh_{N}^{p})^{1/2} | W_{N} - W_{n(k,t)}| \ge \varepsilon, |N - n(k,t)| < \rho n(k,t), \theta = l_{k}\} + \xi/4. \quad (A.12)$$

Fix k with  $1 \le k \le k_0$  and put n(t) = n(k, t). Let  $M_i$  (i=1, 2) be as defined in (A.4) for this n(t). Fix  $t \ge t_3(k) \equiv \max\{t_1(k), t_2\}$ . Then, taking  $M_1 < n(t) \le M_2$  into consideration we get that

$$\begin{split} J(t) &\equiv P\{(Nh_N^p)^{1/2} | W_N - W_{n(t)} | \geq \varepsilon, \ |N - n(t)| < \rho n(t), \ \theta = l_k\} \\ &\leq P\{(ih_i^p)^{1/2} | W_i - W_{n(t)} | \geq \varepsilon \text{ for some } i \text{ with } M_1 < i \leq M_2\} \\ &\leq P\{(\max_{M_1 < i \leq M_2} (ih_i^p)^{1/2}) (\max_{M_1 < i \leq M_2} | W_i - W_{M_1} |) \geq \varepsilon/2\}. \end{split} \tag{A.13}$$

By use of (2.1) and the monotonicity of  $\gamma_n$  we have that for i with  $M_1 < i \le M_2$ 

$$|W_{i} - W_{M_{1}}| \leq \left| \sum_{j=1}^{M_{1}} a_{j} (\beta_{ji} - \beta_{jM_{1}}) Z_{j} \right| + \left| \sum_{j=M_{1}+1}^{i} a_{j} \beta_{ji} Z_{j} \right| + \left| \sum_{j=1}^{i} a_{j} \beta_{ji} \delta_{j} \right| + \left| \sum_{j=1}^{M_{1}} a_{j} \beta_{jM_{1}} \delta_{j} \right|$$

$$\leq (\gamma_{M_{1}} - \gamma_{i}) \left| \sum_{j=1}^{M_{1}} a_{j} \gamma_{j}^{-1} Z_{j} \right| + \gamma_{i} \left| \sum_{j=M_{1}+1}^{i} a_{j} \gamma_{j}^{-1} Z_{j} \right| + 2 \max_{M_{1} \leq i \leq M_{2}} \sum_{j=1}^{i} a_{j} \beta_{ji} \delta_{j} \right|. \quad (A.14)$$

Hence from (A.9), (A.13), (A.14) and the monotonicity of  $nh_n^p$ ,  $h_n^p$  and  $\gamma_n$ 

$$\begin{split} J(t) & \leq P \left\{ (M_2 h_{M_1}^p)^{1/2} (\gamma_{M_1} - \gamma_{M_2}) \left| \sum_{j=1}^{M_1} a_j \gamma_j^{-1} Z_j \right| + (M_2 h_{M_2}^p)^{1/2} \max_{M_1 < i \leq M_2} \gamma_i \left| \sum_{j=M_1+1}^{i} a_j \gamma_j^{-1} Z_j \right| > \varepsilon/4 \right\} \\ & \leq J_1(t) + J_2(t) \,, \end{split} \tag{A.15}$$

where

and

$$J_2(t) \!=\! P\left\{\! (M_2 h_{M_2}^p)^{1/2} \max_{\substack{M_1 \in t \leq M_2}} \! \gamma_i \left| \sum_{j=M_1+1}^i a_j \! \gamma_j^{-1} Z_j \right| > \! \varepsilon/8 \! \right\}.$$

First we shall estimate  $J_1(t)$ . By the Chebychev inequality, (A.5), (A.6) and the assumption of  $EZ_n^2$  we get

$$\begin{split} J_{1}(t) &\leq C_{4} \varepsilon^{-2} (1 - \beta_{M_{1}M_{2}})^{2} (M_{2}/M_{1}) M_{1} h_{M_{1}}^{p} \sum_{j=1}^{M_{1}} a_{j}^{2} \beta_{jM_{1}}^{2} E Z_{j}^{2} \\ &\leq C_{5} \varepsilon^{-2} \{1 - ((1 - \rho)/(1 + \rho))^{a}\}^{2}. \end{split} \tag{A.16}$$

Next we shall estimate  $J_2(t)$ . From the Hájek-Rényi inequality (see Petrov [8], page 51), (A.5) and the monotonicity of  $h_n$  we have

$$\begin{split} J_{2}(t) \leq & C_{6} \varepsilon^{-2} M_{2} h_{M_{2}}^{p} \sum_{j=M_{1}+1}^{M_{2}} a_{j}^{2} E Z_{j}^{2} \leq & C_{7} \varepsilon^{-2} M_{2} h_{M_{2}}^{p} \sum_{j=M_{1}+1}^{M_{2}} j^{-2} h_{j}^{-p} \\ \leq & C_{7} \varepsilon^{-2} M_{2} M_{1}^{-2} (M_{2} - M_{1}) \leq & C_{8} \varepsilon^{-2} \rho . \end{split} \tag{A.17}$$

Set  $C_1 = \max\{C_5, C_8\}$ . Then by (A.2), (A.3), (A.16) and (A.17) we have

$$J_1(t) < \xi/(8k_0)$$
 and  $J_2(t) < \xi/(8k_0)$  for all  $t \ge t_3(k)$ ,

which, together with (A.15), implies that for  $k(1 \le k \le k_0)$ 

$$P\{(Nh_N^p)^{1/2}|W_N - W_{n(k,t)}| \ge \varepsilon, |N - n(k,t)| < \rho n(k,t), \theta = l_k\} < \xi/(4k_0)$$
 for all  $t \ge t_s(k)$ . (A.18)

From (A.12) and (A.18)

$$I_1(t) < \xi/2$$
 for large  $t$ ,

which, together with (A.10) and (A.11), yields that

$$P\{(Nh_N^p)^{1/2}|W_N-W_{[\theta_\pi(t)]}|\geq \varepsilon\} < \xi$$
 for large  $t$ .

Thus the proof of Lemma 3.1 was completed.

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