

## EDGEWORTH EXPANSION FOR ONE-SAMPLE $U$ -STATISTICS

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# EDGEWORTH EXPANSION FOR ONE-SAMPLE *U*-STATISTICS

By

**Yoshihiko MAESONO\***

### Abstract

Under some regularity conditions on kernel, an asymptotic expansion with remainder term  $o(N^{-1})$  is established for one-sample *U*-statistics with kernel of arbitrary degree. This is an extension of the result by Callaert, Janssen and Veraverbeke [1].

### 1. Introduction

Let  $X_1, X_2, \dots, X_N$ , be independently and identically distributed random variables with common distribution function  $F$ . Let  $h$  be symmetric function of its arguments, satisfying  $Eh(X_1, X_2, \dots, X_r)=0$  with  $r \leq N$ ,  $h$  is called a kernel and  $r$  is called its degree. We shall define a one-sample *U*-statistic with a kernel of degree  $r$ ,  $h$ , by

$$U_N = \binom{N}{r}^{-1} \sum_{1 \leq i_1 < i_2 < \dots < i_r \leq N} h(X_{i_1}, X_{i_2}, \dots, X_{i_r}).$$

In the case of degree two, Callaert, Janssen and Veraverbeke [1] have obtained the asymptotic expansion of the distribution of  $U_N$  with the remainder term  $o(N^{-1})$ . In this paper, using the forward martingale characterization of  $U_N$ , we obtain an asymptotic expansion of  $U_N$  with a kernel of arbitrary degree  $r$ . In Section 2 we obtain a representation for  $U_N$  in terms of forward martingales, and get the bounds of absolute moments of martingales. We state the main theorem in Section 3 and prove it in Section 4.

### 2. Preliminaries

We shall represent  $U_N$  in terms of forward martingales. This representation is due to Hoeffding [5] (cf. Serfling [7] p.178). Under the assumption  $E|h(X_1, X_2, \dots, X_r)| < \infty$ , let us define the following notations: for  $1 \leq k \leq r$

$$\begin{aligned} w_k(x_1, x_2, \dots, x_k) &= E\{h(X_1, X_2, \dots, X_r) | X_1 = x_1, X_2 = x_2, \dots, X_k = x_k\}, \\ g_1(x_1) &= w_1(x_1), g_2(x_1, x_2) = w_2(x_1, x_2) - \sum_{i=1}^2 g_1(x_i) \\ &\dots \\ g_r(x_1, x_2, \dots, x_r) &= w_r(x_1, x_2, \dots, x_r) - \sum_{1 \leq i_1 < \dots < i_{r-1} \leq r} g_{r-1}(x_{i_1}, \dots, x_{i_{r-1}}) \\ &\quad - \sum_{1 \leq i_1 < \dots < i_{r-2} \leq r} g_{r-2}(x_{i_1}, \dots, x_{i_{r-2}}) - \dots - \sum_{i=1}^r g_1(x_i), \end{aligned}$$

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and

$$A_{k,N} = \sum_{1 \leq i_1 < i_2 < \dots < i_k \leq N} g_k(X_{i_1}, X_{i_2}, \dots, X_{i_k}), \quad \text{for } 1 \leq k \leq r.$$

Then  $U_N$  can be rewritten as

$$U_N = \binom{N}{r}^{-1} \sum_{k=1}^r \binom{N-k}{r-k} A_{k,N}.$$

It is shown in the proof of Lemma 2 that  $\{A_{k,N}\}_{N \geq k}$  is a forward martingale for each  $k=1, 2, \dots, r$ .

By the definition of  $g_k$ , it is easily shown that if one of  $\{i_1, i_2, \dots, i_k\}$  is not contained in  $\{j_1, j_2, \dots, j_s\}$ , then

$$E\{g_k(X_{i_1}, X_{i_2}, \dots, X_{i_k}) | X_{j_1}, X_{j_2}, \dots, X_{j_s}\} = 0. \tag{2.1}$$

Using this property, we can prove the useful two lemmas.

LEMMA 1. *If  $E|g_k(X_1, X_2, \dots, X_k)| < \infty$  and one of  $\{i_1, i_2, \dots, i_k\}$  is not contained in  $\{j_1, j_2, \dots, j_s\}$ , then for  $f$  satisfying  $E|fg_k| < \infty$ ,*

$$E\{f(X_{j_1}, X_{j_2}, \dots, X_{j_s})g_k(X_{i_1}, X_{i_2}, \dots, X_{i_k})\} = 0.$$

PROOF. Taking the conditional expectation, we have the desired result from (2.1).

Before describing the next lemma, we prepare the notations. For  $1 \leq N_1 < N_2 < \dots < N_k \leq N$  and  $1 \leq k \leq r$ , let us define

$$B_k(N_1, N_2, \dots, N_k) = \sum_{i_1=1}^{N_1} \sum_{i_2=i_1+1}^{N_2} \dots \sum_{i_k=i_{k-1}+1}^{N_k} g_k(X_{i_1}, X_{i_2}, \dots, X_{i_k}).$$

Then we have the upper bound of the  $p$ th absolute moment of  $B_k$ .

LEMMA 2. *Given the existence of the  $p$ th ( $p \geq 2$ ) absolute moment of kernel  $h$ , there exist a positive constant  $C$  such that*

$$E|B_k(N_1, N_2, \dots, N_k)|^p \leq C(\prod_{i=1}^k N_i)^{p/2}. \tag{2.2}$$

*If the second moment of kernel  $h$  is finite, the inequality (2.2) holds also with  $p=1$ .*

PROOF. The latter part of the lemma immediately follows from the former. Therefore we consider the case  $p \geq 2$ . By induction on  $s$  we prove the following inequality for  $1 \leq s \leq k$  and  $1 \leq m_1 < m_2 < \dots < m_s < i_{s+1}, \dots, i_k$ ,

$$\begin{aligned} & E|\sum_{i_1=1}^{m_1} \sum_{i_2=i_1+1}^{m_2} \dots \sum_{i_s=i_{s-1}+1}^{m_s} g_k(X_{i_1}, X_{i_2}, \dots, X_{i_s}, \dots, X_{i_k})|^p \\ & \leq (C_p)^s E|g_k(X_1, X_2, \dots, X_k)|^p (\prod_{i=1}^s m_i)^{p/2} \end{aligned} \tag{2.3}$$

where  $C_p = \{8(p-1) \max(1, 2^{p-3})\}^p$ .

When  $s=1$ , let  $Y_j = \sum_{i_1=1}^{j} g_k(X_{i_1}, X_{i_2}, \dots, X_{i_k})$  for  $j=1, 2, \dots, m_1$ . Then for  $j=1, 2, \dots, m_1$ , we have  $Y_j - Y_{j-1} = g_k(X_j, X_{i_2}, \dots, X_{i_k})$  and  $j < i_2, i_3, \dots, i_k$ , where  $Y_0=0$ . Since  $Y_1, Y_2, \dots, Y_{j-1}$  are functions of  $X_1, X_2, \dots, X_{j-1}, X_{i_2}, \dots, X_{i_k}$ , we find from (2.1) that

$$\begin{aligned} & E\{Y_j - Y_{j-1} | Y_1, Y_2, \dots, Y_{j-1}\} \\ & = E(E\{g_k(X_j, X_{i_2}, \dots, X_{i_k}) | X_1, X_2, \dots, X_{j-1}, X_{i_2}, \dots, X_{i_k}\} | Y_1, Y_2, \dots, Y_{j-1}) = 0. \end{aligned}$$

Therefore  $\{Y_j\}_{0 \leq j \leq m_1}$  is a forward martingale. Applying an upper bound for moments of martingales obtained by Dharmadhikari, Fabian and Jogdeo [2], we have the inequality (2.3) when  $s=1$ .

Assume that (2.3) holds for  $1, 2, \dots, s-1$  instead of  $s$ . For  $s \leq j \leq m_s$ , put  $\tilde{m}_i(j) = \min(m_i, j-s+i)$  (for  $i=1, 2, \dots, s-1$ ) and

$$Z_j = \sum_{i_1=1}^{\tilde{m}_1(j)} \dots \sum_{i_{s-1}=i_{s-2}+1}^{\tilde{m}_{s-1}(j)} \sum_{i_s=i_{s-1}+1}^j g_k(X_{i_1}, \dots, X_{i_{s-1}}, X_{i_s}, \dots, X_{i_k}).$$

Then in the same way of  $s=1$ ,  $\{Z_j\}_{0 \leq j \leq m_s}$  is a forward martingale, where  $Z_j=0$  for  $0 \leq j < s$ . Hence using the result of Dharmadhikari et al. [2], we get the inequality (2.3) for  $s$ . Thus we have the desired result.

### 3. Main Theorem

Before we state the main theorem, we define the following notations:

$$\xi_k^2 = E g_k^2(X_1, X_2, \dots, X_k) \quad \text{for } 1 \leq k \leq r,$$

$$\sigma_N^2 = E\{U_N\}^2 = \frac{r^2}{N} \xi_1^2 + \frac{\{r(r-1)\}^2}{2N(N-1)} \xi_2^2 + \dots + \frac{r!}{N(N-1) \dots (N-r+1)} \xi_r^2,$$

$$\eta(t) = E(\exp\{itg_1(X_1)\}),$$

$$\zeta(x, y) = w_2(x, y) - \frac{r-2}{r-1} \{g_1(x) + g_1(y)\} \quad \text{for } r \geq 2,$$

$$\kappa_3 = \xi_1^{-3} \{E g_1^3(X_1) + 3(r-1)E\{g_1(X_1)g_1(X_2)g_2(X_1, X_2)\}\},$$

$$\begin{aligned} \kappa_4 = & \xi_1^{-4} \{E g_1^4(X_1) - 3\xi_1^4 + 12(r-1)E\{g_1^2(X_1)g_1(X_2)g_2(X_1, X_2)\} \\ & + 12(r-1)^2 E\{g_1(X_2)g_1(X_3)g_2(X_1, X_2)g_2(X_1, X_3)\} \\ & + 4(r-1)(r-2)E\{g_1(X_1)g_1(X_2)g_1(X_3)g_3(X_1, X_2, X_3)\}\}, \end{aligned}$$

and

$$Q_N(x) = \Phi(x) - \phi(x) \left\{ \frac{\kappa_3}{6N^{1/2}}(x^2-1) + \frac{\kappa_4}{24N}(x^3-3x) + \frac{\kappa_5}{72N}(x^5-10x^3+15x) \right\}$$

where  $\Phi(x)$  and  $\phi(x)$  denote the distribution function and the density of the standard normal distribution.

**THEOREM.** *If the following conditions are satisfied (when  $r=1$ , the condition (C) must be omitted)*

- (A)  $E|h(X_1, X_2, \dots, X_r)|^5 < \infty$
- (B)  $\limsup_{|t| \rightarrow \infty} |\eta(t)| < 1$
- (C) *there exist positive constants  $c < 1$  and  $\alpha < 1/8$  such that for  $m = [N^\alpha]$ ,*

$$\begin{aligned} & P\left( \left| E(\exp\{it\sigma_N^{-1} \frac{r(r-1)}{N(N-1)} \sum_{j=m+1}^N \zeta(X_1, X_j)\} | X_{m+1}, \dots, X_N) \right| \leq c \right) \\ & \geq 1 - o\left(\frac{1}{N \log N}\right) \end{aligned}$$

uniformly for all  $t \in [N^{3/4}/\log N, N \log N]$  then

$$\sup_x |P(\sigma_N^{-1} U_N \leq x) - Q_N(x)| = o(N^{-1}).$$

REMARK 1. Instead of condition (A), the asymptotic expansion is valid under the existence of a fourth moment of kernel  $h$ . Lin [6] has proved it in the case of one-sample  $U$ -statistics with kernel of degree two. For arbitrary degree, we can similarly prove it by the way of Lin [6].

REMARK 2. The asymptotic expansion with remainder term  $o(N^{-1/2})$  is valid without condition (C).

**4. Proof of the Theorem**

For  $r=1$ ,  $U_N$  is a sum of independently and identically distributed random variables and the expansion of  $U_N$  has been obtained already (cf. Gnedenko and Kolmogorov [4]). Then we consider the case  $r \geq 2$ .

Let

$$\Psi_N(t) = E \{ \exp(it\sigma_N^{-1}U_N) \}$$

and for  $s=1, 2, 3$ ,

$${}_s\Psi_N(t) = E \{ \exp(it\sum_{k=1}^{s-1}d_{k,N}A_{k,N}) \}$$

where

$$d_{k,N} = \sigma_N^{-1} \binom{N}{r}^{-1} \binom{N-k}{r-k}.$$

Then for  ${}_2\Psi_N(t)$ , we have the following lemma.

LEMMA 3. *If (A) is satisfied, then there exist positive constants  $K_s (s=1, 2, \dots, 6)$  such that for all  $t (-\infty < t < \infty)$ , all integers  $N$  and  $m$  with  $6 < m < N-2$*

$$\begin{aligned} |{}_2\Psi_N(t)| \leq & \left| \eta \left( \frac{rt}{N\sigma_N} \right) \right|^{m-6} (K_1 \sum_{s=0}^3 |t|^s d_{2,N}^s m^s N^s + K_2 |t| d_{3,N} m N^2) \\ & + K_3 t^4 d_{2,N}^4 m^2 N^2 + K_4 t^2 d_{2,N} d_{3,N} m N^{3/2} \\ & + K_5 t^2 d_{3,N}^2 m N^2 + K_6 |t| d_{4,N} m^{1/2} N^{3/2}. \end{aligned} \tag{4.1}$$

Note that if  $r=2$ , the terms which include  $d_{3,N}$  or  $d_{4,N}$  are omitted. Similarly we omit  $K_6 |t| d_{4,N} m^{1/2} N^{3/2}$ , if  $r=3$ .

PROOF. See Appendix 1.

Furthermore, for  ${}_1\Psi_N(t)$  we have Lemma 4.

LEMMA 4. *If (A) is satisfied, then there exist positive constants  $M_s (s=1, 2, \dots, 12)$  such that for all  $t (-\infty < t < \infty)$ , all integers  $N$  and  $m$  with  $8 < m < N-3$*

$$\begin{aligned} |{}_1\Psi_N(t)| \leq & E \{ |E(\exp(itd_{2,N} \sum_{j=m+1}^N \zeta(X_1, X_j)) | X_{m+1}, \dots, X_N) |^{m-8} \} \\ & \times (M_1 \sum_{s=0}^1 |t|^{s+1} d_{2,N}^s d_{4,N} m^{2s+1} N^3 + M_2 \sum_{s=0}^1 |t|^{s+2} d_{2,N}^s d_{3,N}^2 m^{2(s+1)} N^4 \\ & + M_3 \sum_{s=0}^2 |t|^{s+1} d_{2,N}^s d_{3,N} m^{2s+1} N^2 + M_4 \sum_{s=0}^3 |t|^s d_{2,N}^s m^{2s}) \\ & + M_5 |t|^3 d_{2,N}^2 d_{4,N} m^{9/2} N^{3/2} + M_6 t^2 d_{3,N} d_{4,N} m N^{5/2} \\ & + M_7 t^4 d_{2,N}^2 d_{3,N}^2 m^5 N^2 + M_8 t^4 d_{2,N}^2 d_{3,N} m^{13/2} N \\ & + M_9 t^4 d_{2,N}^4 m^8 + M_{10} |t|^3 d_{3,N}^3 m^{3/2} N^3 \\ & + M_{11} t^2 d_{3,N}^2 m N^3 + M_{12} |t| d_{5,N} m^{1/2} N^2. \end{aligned} \tag{4.2}$$

Note that if  $r=2$ , the terms which include  $d_{3,N}$ ,  $d_{4,N}$  or  $d_{5,N}$  are omitted. Similarly if

$r=3$ , we omit the terms which include  $d_{4,N}$  or  $d_{5,N}$ , and if  $r=4$ , omit  $M_{12}|t|d_{5,N}m^{1/2}N^2$ .

PROOF. See Appendix 2.

Now applying Lemma 3, 4 and Esseen's smoothing lemma [3] we shall prove the Theorem. In the sequel, we consider the proof for the case  $r \geq 5$ . When  $r=2, 3, 4$ , we can prove the Theorem more easily.

Let

$$\begin{aligned} \tilde{\Psi}_N(t) &= \int_{-\infty}^{\infty} \exp(itx) dQ_N(x) \\ &= \exp\left(-\frac{t^2}{2}\right) \left\{ 1 + \frac{\kappa_3}{6N^{1/2}}(it)^3 + \frac{\kappa_4}{24N}(it)^4 + \frac{\kappa_5^2}{72N}(it)^6 \right\}. \end{aligned}$$

From smoothing lemma, we have

$$\sup_x |P(\sigma_N^{-1}U_N \leq x) - Q_N(x)| \leq \frac{1}{\pi} \int_{-N \log N}^{N \log N} |t|^{-1} |\Psi_N(t) - \tilde{\Psi}_N(t)| dt + o(N^{-1}).$$

Since the proof for the negative part of  $t$  is similar to that for the positive one, we shall show that

$$\int_0^{N \log N} t^{-1} |\Psi_N(t) - \tilde{\Psi}_N(t)| dt = o(N^{-1}).$$

Since  $d_{k,N} = O(N^{1/2-k})$  and  $E|A_{k,N}| \leq O(N^{k/2})$  from (2.2) in Lemma 2,

$$\begin{aligned} \int_0^{N \log N} t^{-1} |\Psi_N(t) - \tilde{\Psi}_N(t)| dt &\leq \sum_{k=3}^r d_{k,N} E|A_{k,N}| \int_0^{N \log N} dt \\ &= o(N^{-1}). \end{aligned}$$

Similarly we have

$$\int_0^{N^{3/4}/\log N} t^{-1} |{}_1\Psi_N(t) - {}_2\Psi_N(t)| dt = o(N^{-1})$$

and

$$\int_0^{N^{1/4}/\log N} t^{-1} |{}_2\Psi_N(t) - {}_3\Psi_N(t)| dt = o(N^{-1}).$$

Then, putting

$$(I) = \int_0^{N^{1/4}/\log N} t^{-1} |{}_3\Psi_N(t) - \tilde{\Psi}_N(t)| dt,$$

$$(II) = \int_{N^{1/4}/\log N}^{N^{3/4}/\log N} t^{-1} |{}_2\Psi_N(t)| dt,$$

$$(III) = \int_{N^{3/4}/\log N}^{N \log N} t^{-1} |{}_1\Psi_N(t)| dt,$$

and

$$(IV) = \int_{\log N}^{\infty} t^{-1} |\tilde{\Psi}_N(t)| dt,$$

we have

$$\int_0^{N \log N} t^{-1} |\Psi_N(t) - \tilde{\Psi}_N(t)| dt \leq (I) + (II) + (III) + (IV) + o(N^{-1}). \tag{4.3}$$

It immediately follows from condition (A) that (IV) =  $o(N^{-1})$ . Next, we shall prove that (I), (II) and (III) are  $o(N^{-1})$ .

Order of (I): Let us define

$$\begin{aligned} \Psi_N^*(t) = & E\left(\exp(itd_{1,N}A_{1,N})\left\{1+itd_{2,N}A_{2,N}+\frac{(it)^2}{2}d_{2,N}^2A_{2,N}^2\right\}\right) \\ & + E(itd_{3,N}A_{3,N}\exp(itd_{1,N}A_{1,N})). \end{aligned}$$

Then,

$$\begin{aligned} & \int_0^{N^{1/4}/\log N} t^{-1}|\Psi_N^*(t)-\Psi_N(t)| dt \\ & \leq \int_0^{N^{1/4}/\log N} t^{-1}\left|E\left(\exp(itd_{1,N}A_{1,N})\left\{\exp(itd_{2,N}A_{2,N})-\left(1+itd_{2,N}A_{2,N}\right.\right.\right.\right. \\ & \quad \left.\left.\left.+\frac{(it)^2}{2}d_{2,N}^2A_{2,N}^2\right\}\right)\right| dt + \int_0^{N^{1/4}/\log N} t^{-1}|E(itd_{3,N}A_{3,N}\exp(itd_{1,N}A_{1,N})) \\ & \quad \times \{\exp(itd_{2,N}A_{2,N})-1\}| dt + o(N^{-1}). \end{aligned}$$

Using the similar way which has been described in Callaert et al. [1] pp 304-306, we can establish that the first term is  $o(N^{-1})$ . Applying Schwartz's inequality and (2.2) in Lemma 2, we can easily obtain the order of the second term.

Now we evaluate the difference of  $\Psi_N^*(t)$  and  $\tilde{\Psi}_N(t)$ . From Lemma 1  $\Psi_N^*(t)$  can be rewritten as

$$\begin{aligned} \Psi_N^*(t) = & \eta^N(d_{1,N}t) + it\eta^{N-2}(d_{1,N}t)d_{2,N}\frac{N(N-1)}{2}E(\exp\{itd_{1,N}(g_1(X_1)+g_1(X_2))\}g_2(X_1, X_2)) \\ & + \frac{(it)^2}{2}\eta^{N-2}(d_{1,N}t)d_{2,N}^2\frac{N(N-1)}{2}E(\exp\{itd_{1,N}(g_1(X_1)+g_1(X_2))\}g_2^2(X_1, X_2)) \\ & + \frac{(it)^2}{2}\eta^{N-3}(d_{1,N}t)d_{2,N}^2N(N-1)(N-2)E(\exp\{itd_{1,N}(g_1(X_1)+g_1(X_2) \\ & \quad + g_1(X_3))\}g_2(X_1, X_2)g_2(X_1, X_3)) + \frac{(it)^2}{2}\eta^{N-4}(d_{1,N}t)d_{2,N}^2\frac{N(N-1)(N-2)(N-3)}{4} \\ & \times \{E(\exp\{itd_{1,N}(g_1(X_1)+g_1(X_2))\}g_2(X_1, X_2))\}^2 + it\eta^{N-3}(d_{1,N}t)d_{3,N}\frac{N(N-1)(N-2)}{6} \\ & \times E(\exp\{itd_{1,N}(g_1(X_1)+g_1(X_2)+g_1(X_3))\}g_3(X_1, X_2, X_3)). \end{aligned}$$

Let us denote  $\Psi_N^*(t)$  by

$$\begin{aligned} \Psi_N^*(t) = & I_0^* + itd_{2,N}\frac{N(N-1)}{2}I_3^*E_1^* + \frac{(it)^2}{2}d_{2,N}^2\frac{N(N-1)}{2}I_3^*E_2^* \\ & + \frac{(it)^2}{2}d_{2,N}^2N(N-1)(N-2)I_3^*E_3^* + \frac{(it)^2}{2}d_{2,N}^2\frac{N(N-1)(N-2)(N-3)}{4}I_4^*E_4^* \\ & + itd_{3,N}\frac{N(N-1)(N-2)}{6}I_3^*E_5^*. \end{aligned}$$

Then approximations of  $I_k^*$  and  $E_k^*$  are given as follows:

$$\begin{aligned} I_k = & \exp\left(-\frac{t^2}{2}\right)\left(1-\frac{(it)^2}{2N}\left\{\frac{(r-1)^2\xi_2^2}{2\xi_1^2}+k\right\}\right) + \frac{(it)^3}{6N^{1/2}\xi_1^3}Eg_1^3(X_1) \\ & + \frac{(it)^4}{24N\xi_1^4}\{Eg_1^4(X_1)-3\xi_1^4\} + \frac{(it)^6}{72N\xi_1^6}\{Eg_1^6(X_1)\}^2, \end{aligned}$$

$$\begin{aligned}
E_1 &= \frac{(it)^2}{N\xi_1^2} E\{g_1(X_1)g_1(X_2)g_2(X_1, X_2)\} \\
&\quad + \frac{(it)^3}{N^{3/2}\xi_1^3} E\{g_1^2(X_1)g_1(X_2)g_2(X_1, X_2)\}, \\
E_2 &= E g_2^2(X_1, X_2) \\
E_3 &= \frac{(it)^2}{N\xi_1^2} E\{g_1(X_2)g_1(X_3)g_2(X_1, X_2)g_2(X_1, X_3)\}, \\
E_4 &= \left( \frac{(it)^2}{N\xi_1^2} E\{g_1(X_1)g_1(X_2)g_2(X_1, X_2)\} \right)^2, \\
E_5 &= \frac{(it)^3}{N^{3/2}\xi_1^3} E\{g_1(X_1)g_1(X_2)g_1(X_3)g_3(X_1, X_2, X_3)\}.
\end{aligned}$$

By the same way of Lemma 2 as Callaert et al. [1], there exist positive constants  $\varepsilon$  and  $a$  such that for  $0 \leq t \leq \varepsilon N^{1/2}$ ,

$$|I_k^* - I_k| \leq o(N^{-1})P(t) \exp(-at^2)$$

where  $P(t)$  is a polynomial in  $t$ .

Furthermore, from condition (A), and  $\sigma_N^{-1} = N^{1/2}(\xi_1 r)^{-1}(1 + O(N^{-1}))$ , we have the following inequalities:

$$\begin{aligned}
|E_1^* - E_1| &\leq t^4 O(N^{-2}) + t^2 O(N^{-2}) + |t|^3 O(N^{-5/2}), \\
|E_2^* - E_2| &\leq |t| O(N^{-1/2}), \\
|E_3^* - E_3| &\leq |t|^3 O(N^{-3/2}) + t^2 O(N^{-2}), \\
|E_4^* - E_4| &\leq |t|^5 O(N^{-5/2}) + t^6 O(N^{-3}) + t^4 O(N^{-3}), \\
|E_5^* - E_5| &\leq t^4 O(N^{-2}) + |t|^3 O(N^{-5/2}).
\end{aligned}$$

Define

$$\begin{aligned}
\tilde{\Psi}_N^*(t) &= I_0 + itd_{2,N} \frac{N(N-1)}{2} I_2 E_1 + \frac{(it)^2}{2} d_{2,N}^2 \frac{N(N-1)}{2} I_2 E_2 \\
&\quad + \frac{(it)^2}{2} d_{2,N}^2 N(N-1)(N-2) I_3 E_3 + \frac{(it)^2}{2} d_{2,N}^2 \frac{N(N-1)(N-2)(N-3)}{4} I_4 E_4 \\
&\quad + itd_{3,N} \frac{N(N-1)(N-2)}{6} I_3 E_5.
\end{aligned}$$

Then,

$$\begin{aligned}
&\int_0^{N^{1/4}/\log N} t^{-1} |\Psi_N^*(t) - \tilde{\Psi}_N(t)| dt \\
&\leq \int_0^{N^{1/4}/\log N} t^{-1} |\Psi_N^*(t) - \tilde{\Psi}_N^*(t)| dt + \int_0^{N^{1/4}/\log N} t^{-1} |\tilde{\Psi}_N^*(t) - \tilde{\Psi}_N(t)| dt.
\end{aligned}$$

Since  $\sigma_N^{-1} = N^{1/2}(\xi_1 r)^{-1}(1 + O(N^{-1}))$  and for any  $k$ ,

$$\int_0^\infty t^k \exp(-t^2/2) dt < \infty,$$

we get that the last term is  $o(N^{-1})$ .

For the first term, we have



$$\begin{aligned}
|\Psi_N^*(t) - \tilde{\Psi}_N^*(t)| &\leq |I_0^* - I_0| + |t| \frac{r(r-1)}{2\sigma_N} \{|E_1^*(I_2^* - I_2)| + |I_2(E_1^* - E_1)|\} \\
&\quad + \frac{t^2}{2} \frac{\{r(r-1)\}^2}{2\sigma_N N(N-1)} \{|E_2^*(I_3^* - I_3)| + |I_3(E_2^* - E_2)|\} \\
&\quad + \frac{t^2}{2} \frac{\{r(r-1)\}^2(N-2)}{2\sigma_N N(N-1)} \{|E_3^*(I_4^* - I_4)| + |I_4(E_3^* - E_3)|\} \\
&\quad + \frac{t^2}{2} \frac{\{r(r-1)\}^2(N-2)(N-3)}{4\sigma_N N(N-1)} \{|E_4^*(I_5^* - I_5)| + |I_5(E_4^* - E_4)|\} \\
&\quad + |t| \frac{r(r-1)(r-2)}{6\sigma_N} \{|E_5^*(I_6^* - I_6)| + |I_6(E_5^* - E_5)|\}.
\end{aligned}$$

Therefore, using the previous discussions, we can establish that the first term is  $o(N^{-1})$ .

Order of (II): Applying (4.1) in Lemma 3, condition (B) and the same arguments which have been described in Callaert et al. [1] pp 308-309, we get the order of (II).

Order of (III): Combining (4.2) in Lemma 4 and condition (C), we can easily establish that the order of (III) is  $o(N^{-1})$ .

Thus we showed that (I), (II) and (III) are  $o(N^{-1})$  and therefore by (4.3) we have the desired result.

## Appendix

### 1. Proof of Lemma 3

We have

$$\begin{aligned}
{}_2\Psi_N(t) &= E(\exp\{itd_{1,N}B_1(m)\} \exp\{itd_{1,N}(A_{1,N} - B_1(m))\} \exp\{itd_{2,N}B_2(m, N)\} \\
&\quad \times \exp\{itd_{2,N}(A_{2,N} - B_2(m, N))\} \exp\{itd_{3,N}B_3(m, N-1, N)\} \\
&\quad \times \exp\{itd_{3,N}(A_{3,N} - B_3(m, N-1, N))\} \exp\{itd_{4,N}B_4(m, N-2, N-1, N)\} \\
&\quad \times \exp\{itd_{4,N}(A_{4,N} - B_4(m, N-2, N-1, N))\}).
\end{aligned}$$

Let us define  $B_k^* = B_k(m, \dots)$  and  $R_k = \exp\{itd_{k,N}(A_{k,N} - B_k^*)\}$  ( $k=1, 2, 3, 4$ ). Then expanding  $\exp\{itd_{k,N}B_k^*\}$  ( $k=3, 4$ ), we have

$$\begin{aligned}
|{}_2\Psi_N(t)| &\leq |E(\exp\{itd_{1,N}B_1^*\} \exp\{itd_{2,N}B_2^*\} \prod_{j=1}^4 R_j)| \\
&\quad + |t| d_{3,N} |E(\exp\{itd_{1,N}B_1^*\} \prod_{j=1}^3 R_j B_3^*)| \\
&\quad + t^2 d_{2,N} d_{3,N} E|B_2^* B_3^*| + t^2 d_{3,N}^2 E|B_3^*|^2 + |t| d_{4,N} E|B_4^*|.
\end{aligned}$$

Therefore using (2.2) in Lemma 2, we can obtain (4.1) in the same way of Lemma 4 of Callaert et al. [1].

### 2. Proof of Lemma 4

We get

$$\begin{aligned}
{}_1\Psi_N(t) &= E(\exp\{itd_{2,N} \sum_{1 \leq i < j \leq N} \zeta(X_i, X_j)\} \exp\{itd_{3,N}B_3(m, N-1, N)\} \\
&\quad \times \exp\{itd_{3,N}(A_{3,N} - B_3(m, N-1, N))\} \exp\{itd_{4,N}B_4(m, N-2, N-1, N)\}
\end{aligned}$$

$$\begin{aligned} &\times \exp\{itd_{4,N}(A_{4,N} - B_4(m, N-2, N-1, N))\} \\ &\times \exp\{itd_{5,N}B_5(m, N-3, N-2, N-1, N)\} \\ &\times \exp\{itd_{5,N}(A_{5,N} - B_5(m, N-3, N-2, N-1, N))\}. \end{aligned}$$

Let us define  $D = \exp\{itd_{2,N} \sum_{1 \leq i < j \leq N} \zeta(X_i, X_j)\}$ ,  $B_3^* = B_3(m, N-1, N)$ ,  $B_4^* = B_4(m, N-2, N-1, N)$ ,  $B_5^* = B_5(m, N-3, N-2, N-1, N)$  and  $R_k = \exp\{itd_{k,N}(A_{k,N} - B_k^*)\}$  ( $k=3, 4, 5$ ). Then expanding  $\exp\{itd_{k,N}B_k^*\}$  ( $k=3, 4, 5$ ), we have

$$\begin{aligned} |\Psi_N(t)| \leq &\sum_{s=0}^2 |t|^s d_{3,N}^s |E\{DB_3^* \prod_{j=3}^5 R_j\}| + |t|^3 d_{3,N}^3 E|B_3^*|^3 \\ &+ |t|d_{4,N} |E\{DB_4^* \prod_{j=3}^5 R_j\}| + t^2 d_{3,N} d_{4,N} E|B_3^* B_4^*| \\ &+ t^2 d_{4,N}^2 E(B_4^*)^2 + |t|d_{5,N} E|B_5^*| \end{aligned}$$

Using (2.2) in Lemma 2 and applying the factorization of  $D$  which has been discussed in Lemma 5 of Callaert et al. [1], we have the desired result.

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