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EDGEWORTH EXPANSION FOR ONE-SAMPLE *U*-STATISTICS

By

Yoshihiko MAESONO*

Abstract

Under some regularity conditions on kernel, an asymptotic expansion with remainder term $o(N^{-1})$ is established for one-sample U-statistics with kernel of arbitrary degree. This is an extension of the result by Callaert, Janssen and Veraverbeke $\lceil 1 \rceil$.

1. Introduction

Let X_1, X_2, \dots, X_N , be independently and identically distributed random variables with common distribution function F. Let h be symmetric function of its arguments, satisfying $Eh(X_1, X_2, \dots, X_r)=0$ with $r \leq N$, h is called a kernel and r is called its degree. We shall define a one-sample U-statistic with a kernel of degree r, h, by

$$U_{N} = {N \choose r}^{-1} \sum_{1 \le i_{1} < i_{2} < \dots < i_{r} \le N} h(X_{i_{1}}, X_{i_{2}}, \dots, X_{i_{r}}).$$

In the case of degree two, Callaert, Janssen and Veraverbeke [1] have obtained the asymptotic expansion of the distribution of U_N with the remainder term $o(N^{-1})$. In this paper, using the forward martingale characterization of U_N , we obtain an asymptotic expansion of U_N with a kernel of arbitrary degree r. In Section 2 we obtain a representation for U_N in terms of forward martingales, and get the bounds of absolute moments of martingales. We state the main theorem in Section 3 and prove it in Section 4.

2. Preliminaries

We shall represent U_N in terms of forward martingales. This representation is due to Hoeffding [5] (cf. Serfling [7] p. 178). Under the assumption $E \mid h(X_1, X_2, \dots, X_r) \mid < \infty$, let us define the following notations: for $1 \le k \le r$

$$\begin{aligned} w_k(x_1, \ x_2, \ \cdots, \ x_k) &= E\{h(X_1, \ X_2, \ \cdots, \ X_r) | \ X_1 = x_1, \ X_2 = x_2, \ \cdots, \ X_k = x_k\}, \\ g_1(x_1) &= w_1(x_1), \ g_2(x_1, \ x_2) = w_2(x_1, \ x_2) - \sum_{i=1}^2 g_1(x_i) \\ &\cdots \\ g_r(x_1, \ x_2, \ \cdots, \ x_r) &= w_r(x_1, \ x_2, \ \cdots, \ x_r) - \sum_{1 \le i_1 < \cdots < i_{r-1} \le r} g_{r-1}(x_{i_1}, \ \cdots, \ x_{i_{r-1}}) \\ &- \sum_{1 \le i_1 < \cdots < i_{r-2} \le r} g_{r-2}(x_{i_1}, \ \cdots, \ x_{i_{r-2}}) - \cdots - \sum_{i=1}^r g_1(x_i), \end{aligned}$$

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and

$$A_{k,N} = \sum_{1 \le i_1 < i_2 < \dots < i_k \le N} g_k(X_{i_1}, X_{i_2}, \dots, X_{i_k}), \text{ for } 1 \le k \le r.$$

Then U_N can be rewritten as

$$U_N = {N \choose r}^{-1} \sum_{k=1}^{r} {N-k \choose r-k} A_{k,N}.$$

It is shown in the proof of Lemma 2 that $\{A_{k,N}\}_{N\geq k}$ is a forward martingale for each $k=1, 2, \dots, r$.

By the definition of g_k , it is easily shown that if one of $\{i_1, i_2, \dots, i_k\}$ is not contained in $\{j_1, j_2, \dots, j_s\}$, then

$$E\{g_k(X_{i_1}, X_{i_2}, \cdots, X_{i_k}) | X_{j_1}, X_{j_2}, \cdots, X_{j_s}\} = 0.$$
(2.1)

Using this property, we can prove the useful two lemmas.

LEMMA 1. If $E|g_k(X_1, X_2, \dots, X_k)| < \infty$ and one of $\{i_1, i_2, \dots, i_k\}$ is not contained in $\{j_1, j_2, \dots, j_s\}$, then for f satisfying $E|fg_k| < \infty$,

$$E\{f(X_{j_1}, X_{j_2}, \dots, X_{j_s})g_k(X_{i_1}, X_{i_2}, \dots, X_{i_k})\}=0.$$

PROOF. Taking the conditional expectation, we have the desired result from (2.1). Before describing the next lemma, we prepare the notations. For $1 \le N_1 < N_2 < \cdots < N_k \le N$ and $1 \le k \le r$, let us define

$$B_k(N_1, N_2, \cdots, N_k) = \sum_{i_1=1}^{N_1} \sum_{i_2=i_1+1}^{N_2} \cdots \\ \sum_{i_b=i_{b-1}+1}^{N_k} g_k(X_{i_1}, X_{i_2}, \cdots, X_{i_b}).$$

Then we have the upper bound of the pth absolute moment of B_k .

LEMMA 2. Given the existence of the pth $(p \ge 2)$ absolute moment of kernel h, there exist a positive constant C such that

$$E \mid B_k(N_1, N_2, \dots, N_k) \mid {}^p \le C(\prod_{i=1}^k N_i)^{p/2}.$$
 (2.2)

If the second moment of kernel h is finite, the inequality (2.2) holds also with p=1.

PROOF. The latter part of the lemma immediately follows from the former. Therefore we consider the case $p \ge 2$. By induction on s we prove the following inequality for $1 \le s \le k$ and $1 \le m_1 < m_2 < \cdots < m_s < i_{s+1}, \cdots, i_k$,

$$E | \sum_{i=1}^{m_1} \sum_{i_2=i_1+1}^{m_2} \cdots \sum_{i_8=i_{8-1}+1}^{m_8} g_k(X_{i_1}, X_{i_2}, \cdots, X_{i_8}, \cdots, X_{i_k}) |^p$$

$$\leq (C_p)^s E | g_k(X_1, X_2, \cdots, X_k) |^p (\prod_{i=1}^s m_i)^{p/2}$$
(2.3)

where $C_p = \{8(p-1) \max(1, 2^{p-3})\}^p$.

When s=1, let $Y_j = \sum_{i_1=1}^j g_k(X_{i_1}, X_{i_2}, \cdots, X_{i_k})$ for $j=1, 2, \cdots, m_1$. Then for $j=1, 2, \cdots, m_1$, we have $Y_j - Y_{j-1} = g_k(X_j, X_{i_2}, \cdots, X_{i_k})$ and $j < i_2, i_3, \cdots, i_k$, where $Y_0 = 0$. Since $Y_1, Y_2, \cdots, Y_{j-1}$ are functions of $X_1, X_2, \cdots, X_{j-1}, X_{i_2}, \cdots, X_{i_k}$, we find from (2.1) that

$$\begin{split} &E\{Y_{j}-Y_{j-1}|Y_{1},\ Y_{2},\ \cdots,\ Y_{j-1}\}\\ &=E(E\{g_{k}(X_{j},\ X_{i_{2}},\ \cdots,\ X_{i_{k}})|X_{1},\ X_{2},\ \cdots,\ X_{j-1},\ X_{i_{o}},\ \cdots,\ X_{i_{k}}\}|Y_{1},\ Y_{2},\ \cdots,\ Y_{j-1})=0. \end{split}$$

Therefore $\{Y_j\}_{0 \le j \le m_1}$ is a forward martingale. Applying an upper bound for moments of martingales obtained by Dharmadhikari, Fabian and Jogdeo [2], we have the inequality (2.3) when s=1.

Assume that (2.3) holds for 1, 2, \cdots , s-1 instead of s. For $s \leq j \leq m_s$, put $\tilde{m}_i(j) = \min(m_i, j-s+i)$ (for $i=1, 2, \cdots, s-1$) and

$$Z_j = \sum_{i_1=1}^{\tilde{m}_1(j)} \cdots \sum_{i_{s-1}=i_{s-2}+1}^{\tilde{m}_{s-1}(j)} \sum_{i_s=i_{s-1}+1}^j g_k(X_{i_1}, \, \cdots, \, X_{i_{s-1}}, \, X_{i_s}, \, \cdots, \, X_{i_k}).$$

Then in the same way of s=1, $\{Z_j\}_{0 \le j \le m_s}$ is a forward martingale, where $Z_j=0$ for $0 \le j < s$. Hence using the result of Dharmadhikari et al. [2], we get the inequality (2.3) for s. Thus we have the desired result.

3. Main Theorem

Before we state the main theorem, we define the following notations:

$$\begin{split} &\xi_k^2 = E\, g_k^2(X_1,\,X_2,\,\cdots,\,X_k) \qquad \text{for} \quad 1 \leq k \leq r\,, \\ &\sigma_N^2 = E\, \{U_N\}^2 = \frac{r^2}{N}\, \xi_1^2 + \frac{\{r(r-1)\}^2}{2N(N-1)}\, \xi_2^2 + \cdots + \frac{r\,!}{N(N-1)\,\cdots\,(N-r+1)}\, \xi_r^2\,, \\ &\eta(t) = E(\exp\{itg_1(X_1)\})\,, \\ &\zeta(x,\,y) = w_2(x,\,y) - \frac{r-2}{r-1}\, \{g_1(x) + g_1(y)\} \qquad \text{for} \quad r \geq 2\,, \\ &\kappa_3 = \xi_1^{-3}(E\,g_1^3(X_1) + 3(r-1)E\,\{g_1(X_1)g_1(X_2)g_2(X_1,\,X_2)\})\,, \\ &\kappa_4 = \xi_1^{-4}(E\,g_1^4(X_1) - 3\xi_1^4 + 12(r-1)E\,\{g_1^2(X_1)g_1(X_2)g_2(X_1,\,X_2)\} \\ &+ 12(r-1)^2E\,\{g_1(X_2)g_1(X_3)g_2(X_1,\,X_2)g_2(X_1,\,X_3)\} \\ &+ 4(r-1)(r-2)E\,\{g_1(X_1)g_1(X_2)g_1(X_3)g_3(X_1,\,X_2,\,X_3)\})\,, \end{split}$$

and

$$Q_N(x) = \Phi(x) - \phi(x) \left\{ \frac{\kappa_3}{6N^{1/2}} (x^2 - 1) + \frac{\kappa_4}{24N} (x^3 - 3x) + \frac{\kappa_3^2}{72N} (x^5 - 10x^3 + 15x) \right\}$$

where $\Phi(x)$ and $\phi(x)$ denote the distribution function and the density of the standard normal distribution.

THEOREM. If the following conditions are satisfied (when r=1, the condition (C) must be omitted)

- (A) $E | h(X_1, X_2, \dots, X_r) |^5 < \infty$
- (B) $\limsup_{|t|\to\infty} |\eta(t)| < 1$
- (C) there exist positive constants c<1 and $\alpha<1/8$ such that for $m=[N^{\alpha}]$,

$$\begin{split} &P\Big(\left|E(\exp\left\{it\sigma_N^{-1}\frac{r(r-1)}{N(N-1)}\sum_{j=m+1}^N\zeta(X_1,\,X_j)\right\}\right|X_{m+1},\,\cdots,\,X_N)\big|\leq c\Big)\\ &\geq 1-o\Big(\frac{1}{N\log N}\Big) \end{split}$$

uniformly for all $t \in [N^{3/4}/\log N, N \log N]$ then

$$\sup_{x} |P(\sigma_{N}^{-1}U_{N} \leq x) - Q_{N}(x)| = o(N^{-1}).$$

REMARK 1. Instead of condition (A), the asymptotic expansion is valid under the existence of a fourth moment of kernel h. Lin [6] has proved it in the case of one-sample U-statistics with kernel of degree two. For arbitrary degree, we can similarly prove it by the way of Lin [6].

REMARK 2. The asymptotic expansion with remainder term $o(N^{-1/2})$ is valid without condition (C).

4. Proof of the Theorem

For r=1, U_N is a sum of independently and identically distributed random variables and the expansion of U_N has been obtained already (cf. Gnedenko and Kolmogorov [4]). Then we consider the case $r\geq 2$.

Let

$$\Psi_N(t) = E \left\{ \exp\left(it\sigma_N^{-1}U_N\right) \right\}$$

and for s=1, 2, 3,

$$_{s}\Psi_{N}(t) = E \left\{ \exp \left(it \sum_{k=1}^{6-s} d_{k-N} A_{k-N} \right) \right\}$$

where

$$d_{k,N} = \sigma_N^{-1} {N \choose r}^{-1} {N-k \choose r-k}.$$

Then for ${}_{2}\Psi_{N}(t)$, we have the following lemma.

LEMMA 3. If (A) is satisfied, then there exist positive constants $K_s(s=1, 2, \dots, 6)$ such that for all $t(-\infty < t < \infty)$, all integers N and m with 6 < m < N-2

$$|_{2}\Psi_{N}(t)| \leq \left| \eta\left(\frac{rt}{N\sigma_{N}}\right) \right|^{m-6} (K_{1}\sum_{s=0}^{3} |t|^{s} d_{2,N}^{s} m^{s} N^{s} + K_{2} |t| d_{3,N} m N^{2})$$

$$+ K_{3} t^{4} d_{2,N}^{4} m^{2} N^{2} + K_{4} t^{2} d_{2,N} d_{3,N} m N^{3/2}$$

$$+ K_{5} t^{2} d_{3,N}^{2} m N^{2} + K_{6} |t| d_{4,N} m^{1/2} N^{3/2}.$$

$$(4.1)$$

Note that if r=2, the terms which include $d_{3,N}$ or $d_{4,N}$ are omitted. Similarly we omit $K_6|t|d_{4,N}m^{1/2}N^{3/2}$, if r=3.

Proof. See Appendix 1.

Furthermore, for ${}_{1}\Psi_{N}(t)$ we have Lemma 4.

LEMMA 4. If (A) is satisfied, then there exist positive constants $M_s(s=1, 2, \dots, 12)$ such that for all $t(-\infty < t < \infty)$, all integers N and m with 8 < m < N-3

$$\begin{aligned} |_{1} \Psi_{N}(t)| &\leq E(|E(\exp\{itd_{2,N}\sum_{j=m+1}^{N}\zeta(X_{1},X_{j})\}|X_{m+1},\cdots,X_{N})|^{m-8}) \\ &\times (M_{1}\sum_{s=0}^{1}|t|^{s+1}d_{2,N}^{s}d_{4,N}m^{2s+1}N^{3} + M_{2}\sum_{s=0}^{1}|t|^{s+2}d_{2,N}^{s}d_{3,N}^{2}m^{2(s+1)}N^{4} \\ &+ M_{3}\sum_{s=0}^{2}|t|^{s+1}d_{2,N}^{s}d_{3,N}m^{2s+1}N^{2} + M_{4}\sum_{s=0}^{3}|t|^{s}d_{2,N}^{s}m^{2s}) \\ &+ M_{5}|t|^{3}d_{2,N}^{2}d_{4,N}m^{9/2}N^{3/2} + M_{6}t^{2}d_{3,N}d_{4,N}mN^{5/2} \\ &+ M_{7}t^{4}d_{2,N}^{2}d_{3,N}m^{5}N^{2} + M_{8}t^{4}d_{2,N}^{2}d_{3,N}m^{13/2}N \\ &+ M_{9}t^{4}d_{2,N}^{4}m^{8} + M_{10}|t|^{3}d_{3,N}^{3}m^{3/2}N^{3} \\ &+ M_{11}t^{2}d_{4,N}^{2}mN^{3} + M_{12}|t|d_{5,N}m^{1/2}N^{2}. \end{aligned} \tag{4.2}$$

Note that if r=2, the terms which include $d_{3,N}$, $d_{4,N}$ or $d_{5,N}$ are omitted. Similarly if

r=3, we omit the terms which include $d_{4,N}$ or $d_{5,N}$, and if r=4, omit $M_{12}|t|d_{5,N}m^{1/2}N^2$. PROOF. See Appendix 2.

Now applying Lemma 3, 4 and Esseen's smoothing lemma [3] we shall prove the Theorem. In the sequel, we consider the proof for the case $r \ge 5$. When r=2, 3, 4, we can prove the Theorem more easily.

Let

$$\begin{split} \tilde{\varPsi}_N(t) &= \int_{-\infty}^{\infty} \exp{(itx)} dQ_N(x) \\ &= \exp{\left(-\frac{t^2}{2}\right)} \Big\{ 1 + \frac{\kappa_3}{6N^{1/2}} (it)^3 + \frac{\kappa_4}{24N} (it)^4 + \frac{\kappa_3^2}{72N} (it)^6 \Big\} \,. \end{split}$$

From smoothing lemma, we have

$$\sup_{x} |P(\sigma_{N}^{-1}U_{N} \leq x) - Q_{N}(x)| \leq \frac{1}{\pi} \int_{-N \log N}^{N \log N} |t|^{-1} |\Psi_{N}(t) - \tilde{\Psi}_{N}(t)| dt + o(N^{-1}).$$

Since the proof for the negative part of t is similar to that for the positive one, we shall show that

$$\int_{0}^{N \log N} t^{-1} | \Psi_{N}(t) - \tilde{\Psi}_{N}(t) | dt = o(N^{-1}).$$

Since $d_{k,N} = O(N^{1/2-k})$ and $E|A_{k,N}| \leq O(N^{k/2})$ from (2.2) in Lemma 2,

$$\int_{0}^{N \log N} t^{-1} | \Psi_{N}(t) - {}_{1}\Psi_{N}(t) | dt \leq \sum_{k=0}^{r} d_{k,N} E | A_{k,N} | \int_{0}^{N \log N} dt$$

$$= o(N^{-1}).$$

Similary we have

and

$$\int_0^{N^{3/4/\log N}} t^{-1} |_1 \Psi_N(t) - _2 \Psi_N(t) | dt = o(N^{-1})$$

$$\int_0^{N^{1/4/\log N}} t^{-1} |_2 \Psi_N(t) - _3 \Psi_N(t) | dt = o(N^{-1}) .$$

Then, putting

$$\begin{split} (\text{ I }) &= \int_{0}^{N^{1/4}/\log N} t^{-1} |_{3} \varPsi_{N}(t) - \tilde{\varPsi}_{N}(t) | dt \,, \\ (\text{ II }) &= \int_{N^{1/4}/\log N}^{N^{3/4}/\log N} t^{-1} |_{2} \varPsi_{N}(t) | dt \,, \\ (\text{ III }) &= \int_{N^{3/4}/\log N}^{N\log N} t^{-1} |_{1} \varPsi_{N}(t) | dt \,, \end{split}$$

and

$$(\text{IV}) = \int_{\log N}^{\infty} t^{-1} | \widetilde{\Psi}_N(t) | dt,$$

we have

$$\int_{0}^{N \log N} t^{-1} | \Psi_{N}(t) - \tilde{\Psi}_{N}(t) | dt \leq (I) + (II) + (III) + (IV) + o(N^{-1}).$$
 (4.3)

It immediately follows from condition (A) that (IV)= $o(N^{-1})$. Next, we shall prove that (I), (II) and (III) are $o(N^{-1})$.

Order of (I): Let us define

$$\Psi_{N}^{*}(t) = E\left(\exp\left(itd_{1,N}A_{1,N}\right)\left\{1 + itd_{2,N}A_{2,N} + \frac{(it)^{2}}{2}d_{2,N}^{2}A_{2,N}^{2}\right\}\right) + E\left(itd_{3,N}A_{3,N}\exp\left(itd_{1,N}A_{1,N}\right)\right).$$

Then.

$$\begin{split} & \int_{0}^{N^{1/4}/\log N} t^{-1} | \varPsi_{N}^{*}(t) - {}_{3}\varPsi_{N}(t) | \, dt \\ & \leq & \int_{0}^{N^{1/4}/\log N} t^{-1} \Big| E\Big(\exp(itd_{1,\,N}A_{1,\,N}) \Big\{ \exp(itd_{2,\,N}A_{2,\,N}) - \Big(1 + itd_{2,\,N}A_{2,\,N} \\ & + \frac{(it)^{2}}{2} d_{2,\,N}^{2} A_{2,\,N}^{2} \Big) \Big\} \Big) \Big| \, dt + \int_{0}^{N^{1/4}/\log N} t^{-1} | \, E(itd_{3,\,N}A_{3,\,N} \exp(itd_{1,\,N}A_{1,\,N}) \\ & \times \{ \exp(itd_{2,\,N}A_{2,\,N}) - 1 \}) | \, dt + o(N^{-1}) \, . \end{split}$$

Using the similar way which has been described in Callaert et al. [1] pp 304-306, we can establish that the first term is $o(N^{-1})$. Applying Schwartz's inequality and (2.2) in Lemma 2, we can easily obtain the order of the second term.

Now we evaluate the difference of $\Psi_N^*(t)$ and $\tilde{\Psi}_N(t)$. From Lemma 1 $\Psi_N^*(t)$ can be rewritten as

$$\begin{split} \varPsi_N^*(t) &= \eta^N(d_{1,N}t) + it\eta^{N-2}(d_{1,N}t)d_{2,N}\frac{N(N-1)}{2}E(\exp\{itd_{1,N}(g_1(X_1) + g_1(X_2))\}g_2(X_1,X_2)) \\ &+ \frac{(it)^2}{2}\eta^{N-2}(d_{1,N}t)d_{2,N}^2\frac{N(N-1)}{2}E(\exp\{itd_{1,N}(g_1(X_1) + g_1(X_2))\}g_2^2(X_1,X_2)) \\ &+ \frac{(it)^2}{2}\eta^{N-3}(d_{1,N}t)d_{2,N}^2N(N-1)(N-2)E(\exp\{itd_{1,N}(g_1(X_1) + g_1(X_2))\}g_2(X_1,X_2)) \\ &+ g_1(X_3))\}g_2(X_1,X_2)g_2(X_1,X_3)) + \frac{(it)^2}{2}\eta^{N-4}(d_{1,N}t)d_{2,N}^2\frac{N(N-1)(N-2)(N-3)}{4} \\ &\times \{E(\exp\{itd_{1,N}(g_1(X_1) + g_1(X_2))\}g_2(X_1,X_2))\}^2 + it\eta^{N-3}(d_{1,N}t)d_{3,N}\frac{N(N-1)(N-2)}{6} \\ &\times E(\exp\{itd_{1,N}(g_1(X_1) + g_1(X_2) + g_1(X_3))\}g_3(X_1,X_2,X_3))\,. \end{split}$$

Let us denote $\Psi_N^*(t)$ by

$$\begin{split} \varPsi_{N}^{*}(t) &= I_{0}^{*} + itd_{2,N} \frac{N(N-1)}{2} I_{2}^{*} E_{1}^{*} + \frac{(it)^{2}}{2} d_{2,N}^{2} \frac{N(N-1)}{2} I_{2}^{*} E_{2}^{*} \\ &+ \frac{(it)^{2}}{2} d_{2,N}^{2} N(N-1)(N-2) I_{3}^{*} E_{3}^{*} + \frac{(it)^{2}}{2} d_{2,N}^{2} \frac{N(N-1)(N-2)(N-3)}{4} I_{4}^{*} E_{4}^{*} \\ &+ itd_{3,N} \frac{N(N-1)(N-2)}{6} I_{3}^{*} E_{5}^{*}. \end{split}$$

Then approximations of I_k^* and E_k^* are given as follows:

$$\begin{split} I_{k} = & \exp\left(-\frac{t^{2}}{2}\right) \left(1 - \frac{(it)^{2}}{2N} \left\{\frac{(r-1)^{2}\xi_{2}^{2}}{2\xi_{1}^{2}} + k\right\} + \frac{(it)^{3}}{6N^{1/2}\xi_{1}^{3}} Eg_{1}^{3}(X_{1}) + \frac{(it)^{4}}{24N\xi_{1}^{4}} \left\{Eg_{1}^{4}(X_{1}) - 3\xi_{1}^{4}\right\} + \frac{(it)^{6}}{72N\xi_{1}^{5}} \left\{Eg_{1}^{3}(X_{1})\right\}^{2}\right), \end{split}$$

$$\begin{split} E_1 &= \frac{(it)^2}{N\xi_1^2} E\{g_1(X_1)g_1(X_2)g_2(X_1, X_2)\} \\ &+ \frac{(it)^3}{N^{3/2}\xi_1^3} E\{g_1^2(X_1)g_1(X_2)g_2(X_1, X_2)\}, \\ E_2 &= Eg_2^2(X_1, X_2) \\ E_3 &= \frac{(it)^2}{N\xi_1^2} E\{g_1(X_2)g_1(X_3)g_2(X_1, X_2)g_2(X_1, X_3)\}, \\ E_4 &= \left(\frac{(it)^2}{N\xi_1^2} E\{g_1(X_1)g_1(X_2)g_2(X_1, X_2)\}\right)^2, \\ E_5 &= \frac{(it)^3}{N^{3/2}\xi_1^3} E\{g_1(X_1)g_1(X_2)g_1(X_3)g_3(X_1, X_2, X_3)\}. \end{split}$$

By the same way of Lemma 2 as Callaert et al. [1], there exist positive constants ε and a such that for $0 \le t \le \varepsilon N^{1/2}$,

$$|I_k^* - I_k| \le o(N^{-1})P(t) \exp(-at^2)$$

where P(t) is a polynomial in t.

Furthermore, from condition (A), and $\sigma_N^{-1} = N^{1/2} (\xi_1 r)^{-1} (1 + O(N^{-1}))$, we have the following inequalities:

$$\begin{split} |E_1^* - E_1| &\leq t^4 O(N^{-2}) + t^2 O(N^{-2}) + |t|^3 O(N^{-5/2}), \\ |E_2^* - E_2| &\leq |t| O(N^{-1/2}), \\ |E_3^* - E_3| &\leq |t|^3 O(N^{-3/2}) + t^2 O(N^{-2}), \\ |E_4^* - E_4| &\leq |t|^5 O(N^{-5/2}) + t^6 O(N^{-3}) + t^4 O(N^{-3}), \\ |E_5^* - E_5| &\leq t^4 O(N^{-2}) + |t|^3 O(N^{-5/2}). \end{split}$$

Define

$$\begin{split} \tilde{\mathscr{Y}}_{N}^{*}(t) &= I_{0} + itd_{2, N} \frac{N(N-1)}{2} I_{2}E_{1} + \frac{(it)^{2}}{2} d_{2, N}^{2} \frac{N(N-1)}{2} I_{2}E_{2} \\ &+ \frac{(it)^{2}}{2} d_{2, N}^{2} N(N-1)(N-2) I_{3}E_{3} + \frac{(it)^{2}}{2} d_{2, N}^{2} \frac{N(N-1)(N-2)(N-3)}{4} I_{4}E_{4} \\ &+ itd_{3, N} \frac{N(N-1)(N-2)}{6} I_{3}E_{5}. \end{split}$$

Then,

$$\begin{split} & \int_0^{N^{1/4/\log N}} t^{-1} | \varPsi_N^*(t) - \tilde{\varPsi}_N(t) | \, dt \\ & \leq & \int_0^{N^{1/4/\log N}} t^{-1} | \varPsi_N^*(t) - \tilde{\varPsi}_N^*(t) | \, dt + \int_0^{N^{1/4/\log N}} t^{-1} | \tilde{\varPsi}_N^*(t) - \tilde{\varPsi}_N(t) | \, dt \, . \end{split}$$

Since $\sigma_N^{-1} = N^{1/2}(\xi_1 r)^{-1}(1 + O(N^{-1}))$ and for any k,

$$\int_0^\infty t^k \exp(-t^2/2) dt < \infty,$$

we get that the last term is $o(N^{-1})$.

For the first term, we have

$$\begin{split} |\Psi_N^*(t) - \tilde{\Psi}_N^*(t)| &\leq |I_0^* - I_0| + |t| \frac{r(r-1)}{2\sigma_N} \left\{ |E_1^*(I_2^* - I_2)| + |I_2(E_1^* - E_1)| \right\} \\ &+ \frac{t^2}{2} \frac{\{r(r-1)\}^2}{2\sigma_N N(N-1)} \left\{ |E_2^*(I_2^* - I_2)| + |I_2(E_2^* - E_2)| \right\} \\ &+ \frac{t^2}{2} \frac{\{r(r-1)\}^2(N-2)}{\{2\sigma_N N(N-1)} \left\{ |E_3^*(I_3^* - I_3)| + |I_3(E_3^* - E_3)| \right\} \\ &+ \frac{t^2}{2} \frac{\{r(r-1)\}^2(N-2)(N-3)}{4\sigma_N N(N-1)} \left\{ |E_4^*(I_4^* - I_4)| + |I_4(E_4^* - E_4)| \right\} \\ &+ |t| \frac{r(r-1)(r-2)}{6\sigma_N} \left\{ |E_5^*(I_3^* - I_3)| + |I_3(E_5^* - E_5)| \right\}. \end{split}$$

Therefore, using the previous discussions, we can establish that the first term is $o(N^{-1})$.

Order of (II): Applying (4.1) in Lemma 3, condition (B) and the same arguments which have been described in Callaert et al. [1] pp 308-309, we get the order of (II).

Order of (III): Combining (4.2) in Lemma 4 and condition (C), we can easily establish that the order of (III) is $o(N^{-1})$.

Thus we showed that (I), (II) and (III) are $o(N^{-1})$ and therefore by (4.3) we have the desired result.

Appendix

1. Proof of Lemma 3

We have

$$\begin{split} {}_{2}\Psi_{N}(t) &= E(\exp\{itd_{1,N}B_{1}(m)\}\exp\{itd_{1,N}(A_{1,N} - B_{1}(m))\}\exp\{itd_{2,N}B_{2}(m,N)\}\\ &\times \exp\{itd_{2,N}(A_{2,N} - B_{2}(m,N))\}\exp\{itd_{3,N}B_{3}(m,N-1,N)\}\\ &\times \exp\{itd_{3,N}(A_{3,N} - B_{3}(m,N-1,N))\}\exp\{itd_{4,N}B_{4}(m,N-2,N-1,N)\}\\ &\times \exp\{itd_{4,N}(A_{4,N} - B_{4}(m,N-2,N-1,N))\}). \end{split}$$

Let us define $B_k^* = B_k(m, \dots)$ and $R_k = \exp\{itd_{k,N}(A_{k,N} - B_k^*)\}$ (k=1, 2, 3, 4). Then expanding $\exp\{itd_{k,N}B_k^*\}$ (k=3, 4), we have

$$\begin{split} |_{2} \Psi_{N}(t)| & \leq |E(\exp\{itd_{1,N}B_{1}^{*}\}\exp\{itd_{2,N}B_{2}^{*}\}\prod_{j=1}^{4}R_{j})| \\ & + |t|d_{3,N}|E(\exp\{itd_{1,N}B_{1}^{*}\}\prod_{j=1}^{4}R_{j}B_{3}^{*})| \\ & + t^{2}d_{2,N}d_{3,N}E|B_{2}^{*}B_{3}^{*}| + t^{2}d_{3,N}^{2}E|B_{3}^{*}|^{2} + |t|d_{4,N}E|B_{4}^{*}|. \end{split}$$

Therefore using (2.2) in Lemma 2, we can obtain (4.1) in the same way of Lemma 4 of Callaert et al. [1].

2. Proof of Lemma 4

We get

$${}_{1}\Psi_{N}(t) = E(\exp\{itd_{2,N}\sum_{1 \le i < j \le N}\zeta(X_{i}, X_{j})\}\exp\{itd_{3,N}B_{3}(m, N-1, N)\}$$

$$\times \exp\{itd_{3,N}(A_{3,N}-B_{3}(m, N-1, N))\}\exp\{itd_{4,N}B_{4}(m, N-2, N-1, N)\}$$

$$\times \exp\{itd_{4,N}(A_{4,N}-B_4(m, N-2, N-1, N))\}\$$
 $\times \exp\{itd_{5,N}B_5(m, N-3, N-2, N-1, N)\}\$
 $\times \exp\{itd_{5,N}(A_{5,N}-B_5(m, N-3, N-2, N-1, N))\}).$

Let us define $D=\exp\{itd_{2,N}\sum_{1\leq i< j\leq N}\zeta(X_i,X_j)\}$, $B_3^*=B_3(m,N-1,N)$, $B_4^*=B_4(m,N-2,N-1,N)$, $B_5^*=B_5(m,N-3,N-2,N-1,N)$ and $R_k=\exp\{itd_{k,N}(A_{k,N}-B_k^*)\}$ (k=3,4,5). Then expanding $\exp\{itd_{k,N}B_k^*\}$ (k=3,4,5), we have

$$\begin{aligned} | {}_{1} \varPsi_{N}(t) | & \leq \sum_{s=0}^{2} |t|^{s} d_{3,N}^{s} | E\{DB_{3}^{*s} \Pi_{j=3}^{5} R_{j}\} | + |t|^{3} d_{3,N}^{3} E |B_{3}^{*}|^{3} \\ & + |t| d_{4,N} | E\{DB_{4}^{*} \Pi_{j=3}^{5} R_{j}\} | + t^{2} d_{3,N} d_{4,N} E |B_{3}^{*} B_{4}^{*}| \\ & + t^{2} d_{3,N}^{2} E(B_{4}^{*})^{2} + |t| d_{5,N} E |B_{5}^{*}| \end{aligned}$$

Using (2.2) in Lemma 2 and applying the factorization of D which has been discussed in Lemma 5 of Callaert et al. [1], we have the desired result.

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