

## STRONG CONSISTENCY OF A SEQUENTIAL ESTIMATOR OF A PROBABILITY DENSITY FUNCTION

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# STRONG CONSISTENCY OF A SEQUENTIAL ESTIMATOR OF A PROBABILITY DENSITY FUNCTION

By

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## 1. Introduction.

Consider a sequence  $X_1, X_2, X_3 \dots$  of independent identically distributed  $m$ -dimensional random vectors having a probability density function,  $f(x)$ . Van Ryzin (1969) has shown that under appropriate conditions, estimates of the form

$$(1) \quad f_n(x) = n^{-1} \sum_{j=1}^n K_n(x, X_j),$$

where

$$K_n(x, X_j) = h_n^{-m} K(h_n^{-1}(x - X_j)),$$

are strongly consistent. That is, if  $x$  is contained in the continuity set  $C(f)$ , then  $f_n(x) \rightarrow f(x)$  with probability one. Here  $K(u) = K(u_1, u_2, \dots, u_m)$  is a real valued, Borel measurable function on  $R^m$ , where  $R^m$  is the  $m$ -fold convolution of the real line, such that

$$(2) \quad K(u) \text{ is a density on } R^m$$

$$(3) \quad \sup_{u \in R^m} |K(u)| < \infty$$

$$(4) \quad \|u\| K(u) \rightarrow 0 \quad \text{as} \quad \|u\|^2 = \sum_{i=1}^m u_i^2 \rightarrow \infty$$

and  $\{h_n\}$  is a sequence of numbers such that

$$(5) \quad h_n > 0, \quad n = 1, 2, \dots; \quad \lim_{n \rightarrow \infty} h_n = 0, \quad \text{and} \quad \lim_{n \rightarrow \infty} n h_n = \infty.$$

Yamato, (1970) considers a modified estimator of the form

$$(6) \quad \hat{f}(x) = n^{-1} \sum_{j=1}^n h_j^{-m} K(h_j^{-m}(x - X_j))$$

and shows the weak consistency of this estimator. (That is  $\lim_{n \rightarrow \infty} E|\hat{f}_n(x) - f(x)|^2 = 0$  at all points  $x$ ). It is our purpose to show, using Van Ryzin's method that this modified estimator is also strongly consistent. However, because of the special form that this estimator takes, Van Ryzin's conditions can be somewhat weakened.

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## 2. Strong Consistency.

Pointwise: In order to show strong consistency we will use the following lemma which is proved in Van Ryzin (1969).

LEMMA 1. Let  $\{Y_n\}$  and  $\{Y'_n\}$  be two sequences of random variables on a probability space  $(\Omega, F, P)$ . Let  $\{F_n\}$  be a sequence of Borel fields,  $F_n \subset F_{n+1} \subset F$ , where  $Y_n$  and  $Y'_n$  are measurable with respect to  $F_n$ . If

- (i)  $0 \leq Y_n$  a. e.
- (ii)  $EY_1 < \infty$
- (iii)  $E[Y_{n+1}|F_n] \leq Y_n + Y'_n$  a. e.
- (iv)  $\sum_{n=1}^{\infty} E|Y'_n| < \infty$

then  $Y_n$  converges a. e. to a finite limit.

THEOREM 1. If  $K(u)$  satisfies (2)–(4) and  $\{h_n\}$  is a monotone decreasing sequence of numbers satisfying (5) and if in addition

$$(7) \quad \sum_{n=1}^{\infty} n^{-2} h_n^{-m} < \infty$$

$$(8) \quad \lim_{n \rightarrow \infty} n^{-1} \sum_{j=1}^n h_j^{-m} h_n^m = \alpha, \quad 0 \leq \alpha \leq 1$$

then  $\hat{f}_n(x) \rightarrow f(x)$  with probability one if  $x \in C(f)$ .

PROOF.  $E\hat{f}_n(x) \rightarrow f(x)$  as  $n \rightarrow \infty$  (Yamato (1970)). Thus, it is sufficient to show  $\hat{f}_n(x) - E\hat{f}_n(x) \rightarrow 0$  with probability one as  $n \rightarrow \infty$ .

Let  $Y_n = \{\hat{f}_n(x) - E\hat{f}_n(x)\}^2$  and  $F_n =$  Borel field generated by  $X_1, X_2, \dots, X_n$  and note that  $\lim_{n \rightarrow \infty} E(Y_n) = 0$  (Yamato (1970)). Since,

$$(9) \quad \hat{f}_n(x) = n^{-1}(n-1)\hat{f}_{n-1}(x) + n^{-1}h_n^{-m}K(h_n^{-1}(x - X_n)),$$

it follows that

$$\begin{aligned} \hat{f}_{n+1}(x) - E\hat{f}_{n+1}(x) &= \hat{f}_n(x) - E\hat{f}_n(x) - (n+1)^{-1}\{\hat{f}_n(x) - E\hat{f}_n(x)\} \\ &\quad + (n+1)^{-1}\{h_{n+1}^{-m}K(h_{n+1}^{-1}(x - X_{n+1})) - h_{n+1}^{-m}E(K(h_{n+1}^{-1}(x - X_{n+1})))\}. \end{aligned}$$

Thus,

$$E[Y_{n+1}|F_n] = Y_n + (n+1)^{-2}Y_n - 2(n+1)^{-1}Y_n + (n+1)^{-2}h_{n+1}^{-2m} \text{Var } K(h_{n+1}^{-1}(x - X)).$$

Let  $Y'_n = (n+1)^{-2}Y_n - 2(n+1)^{-1}Y_n + (n+1)^{-2}h_{n+1}^{-2m} \text{Var } K(h_{n+1}^{-1}(x - X))$ .

In order to use Lemma 1 we must verify

$$(10) \quad \sum_{n=1}^{\infty} E|Y'_n| < \infty.$$

But,

$$\sum_{n=1}^{\infty} (n+1)^{-1}E|Y_n| = \sum_{n=1}^{\infty} (n+1)^{-1} \text{Var } \hat{f}_n(x).$$

Now, Yamato (1970), using a monotone sequence satisfying (5) and (8) has shown (Theorem 3) that

$$(11) \quad \lim_{n \rightarrow \infty} nh_n^m \text{Var } \hat{f}_n(x) = \alpha f(x) \int_{R^m} K^2(y) dy.$$

Thus  $n^{-1} \text{Var } \hat{f}_n(x) = O(n^{-2} h_n^{-m})$  for large  $n$ . Hence,  $\sum_{n=1}^{\infty} (n+1)^{-1} \text{Var } \hat{f}_n(x) < \infty$  using condition (7).

The fact that  $\sum_{n=1}^{\infty} (n+1)^{-2} h_{n+1}^{-2m} \text{Var } K(h_{n+1}^{-1}(x-X)) < \infty$  is shown to follow from condition (7) by Van Ryzin (1969).

Thus  $\sum_{n=1}^{\infty} E|Y'_n| < \infty$  and the result follows by using Lemma 1 and the fact that the mean square limit and the almost sure limit coincide with probability one.

Uniform:

Let  $k(t) = k(t_1, \dots, t_m) = \int e^{it'u} K(u) du$  where  $t'u = \sum_{j=1}^m t_j u_j$ . Also let  $\phi_n(t) = n^{-1} \sum_{j=1}^n e^{it'x_j}$  and let  $\phi(t) = E e^{it'x}$ .

THEOREM 2. If  $K(u)$  satisfies conditions (2)–(4) and  $\{h_n\}$  is a monotone decreasing sequence of real numbers satisfying (5) and if both  $h_n h_{n+1}^{-1} \rightarrow 1$  and  $n h_n^{-2m} \rightarrow \infty$  as  $n \rightarrow \infty$ , and if

$$(10) \quad \sum_{n=1}^{\infty} (n h_n^m)^{-2} < \infty$$

$$(11) \quad \sum_{n=1}^{\infty} \frac{1}{n h_n^{2m-1}} \left| \frac{1}{h_{n+1}} - \frac{1}{h_n} \right| < \infty$$

and if also

$$(12) \quad \int |k(t)| dt < \infty$$

where  $|k(u)|$  is non-decreasing on  $u < 0$  and non-increasing on  $u \geq 0$ , then if  $f(x)$  is uniformly continuous on  $R^m$ ,  $\sup_x |\hat{f}_n(x) - f(x)| \rightarrow 0$  with probability one as  $n \rightarrow \infty$ .

PROOF. (All integrals are over  $R^m$ , unless otherwise specified.) We first prove  $\sup_x |\hat{f}_n(x) - E \hat{f}_n(x)| \rightarrow 0$  a.s. as  $n \rightarrow \infty$ . Since both  $k(u)$  and  $K(u)$  are in  $L_1$ , using the inversion theorem for a Fourier transform

$$\begin{aligned} \sup_x |\hat{f}_n(x) - E \hat{f}_n(x)| &= \sup_x \left| (2\pi)^{-1} \int n^{-1} \sum_{j=1}^n [e^{i u' x_j} - \phi(u)] k(h_j u) e^{-i u' x} du \right| \\ &\leq (2\pi)^{-m} \int \left| n^{-1} \sum_{j=1}^n [e^{i u' x_j} - \phi(u)] k(h_j u) \right| du \\ &= (2\pi)^{-m} \int \left| n^{-1} \sum_{j=1}^n [e^{i u' x_j} - \phi(u)] \frac{k(h_j u)}{k(h_n u)} \right| |k(h_n u)| du. \end{aligned}$$

Using Schwartz's inequality, this is less than

$$\begin{aligned} &(2\pi)^{-m} \left\{ \int |k(h_n u)| du \right\}^{\frac{1}{2}} \left\{ \int \left| n^{-1} \sum_{j=1}^n [e^{i u' x_j} - \phi(u)] \frac{k(h_j u)}{k(h_n u)} \right|^2 |k(h_n u)| du \right\}^{\frac{1}{2}} \\ &= (2\pi)^{-m} \left\{ \int h_n^{-m} |k(u)| du \right\}^{\frac{1}{2}} \left\{ \int \left| n^{-1} \sum_{j=1}^n [e^{i u' x_j} - \phi(u)] \frac{k(h_j u)}{k(h_n u)} \right|^2 |k(h_n u)| du \right\}^{\frac{1}{2}}. \end{aligned}$$

Since  $\int |k(u)| du < \infty$  by assumption (12), we need only consider the term

$$\int h_n^{-m} \left| n^{-1} \sum_{j=1}^n [e^{i u' x_j} - \phi(u)] \frac{k(h_j u)}{k(h_n u)} \right|^2 |k(h_n u)| du$$

which we will denote by  $Y_n$ . Taking expectations we get

$$EY_n = \int h_n^{-m} E |n^{-1} \sum_{j=1}^n [e^{iu'X_j} - \phi(u)] \frac{k(h_j u)}{k(h_n u)}|^2 |k(h_n u)| du.$$

But,

$$E |e^{iu'X_j} - \phi(u)|^2 = E [e^{iu'X_j} - \phi(u)] [e^{-iu'X_j} - \overline{\phi(u)}] = 1 - |\phi(u)|^2$$

so that, we obtain

$$(13) \quad EY_n = n^{-2} h_n^{-m} \int (1 - |\phi(u)|^2) \sum_{j=1}^n \left| \frac{k(h_j u)}{k(h_n u)} \right|^2 |k(h_n u)| du.$$

However, since  $|k(u)|$  is non-decreasing on  $u < 0$  and non-increasing on  $u \geq 0$ , we have

$$(14) \quad \left| \frac{k(h_j u)}{k(h_n u)} \right| \leq 1, \quad j \leq n.$$

Thus, using inequality (14) in equation (13) we get

$$EY_n \leq n^{-1} h_n^{-m} \int |k(h_n u)| du = n^{-1} h_n^{-2m} \int |k(u)| du.$$

Condition (12) and the assumption  $nh_n^{2m} \rightarrow \infty$  ensure that  $EY_n$  tends to zero as  $n \rightarrow \infty$ . We now prove  $Y_n$  converges with probability one. Write,

$$\begin{aligned} Y_{n+1} &= h_{n+1}^{-m} \int (n+1)^{-2} \left| \sum_{j=1}^{n+1} [e^{iu'X_j} - \phi(u)] \frac{k(h_j u)}{k(h_{n+1} u)} \right|^2 |k(h_{n+1} u)| du \\ &= (n+1)^{-2} h_{n+1}^{-m} \int \left| \sum_{j=1}^n [e^{iu'X_j} - \phi(u)] \frac{k(h_j u)}{k(h_n u)} + e^{iu'X_{n+1}} - \phi(u) \right|^2 |k(h_{n+1} u)| du. \end{aligned}$$

Let  $F_n$  be defined as in Theorem 1. We now expand the first term in the integral using the fact that something of the form  $|z|^2 = z\bar{z}$  and that

$$E(e^{iu'X_{n+1}} - \phi(u) | F_n) = 0 \quad \text{and} \quad E(e^{-iu'X_{n+1}} - \overline{\phi(u)} | F_n) = 0.$$

Thus,

$$\begin{aligned} E[Y_{n+1} | F_n] &= (n+1)^{-2} h_{n+1}^{-m} \int \left| \sum_{j=1}^n [e^{iu'X_j} - \phi(u)] \frac{k(h_j u)}{k(h_n u)} \right|^2 |k(h_{n+1} u)| du \\ &\quad + h_{n+1}^{-m} (n+1)^{-2} \int (1 - |\phi(u)|^2) |k(h_{n+1} u)| du. \end{aligned}$$

But  $(1 - |\phi(u)|^2) \leq 1$  and  $|k(u)| \leq 1$  since we assume  $K(u)$  is a density function. Thus,

$$\begin{aligned} E[Y_{n+1} | F_n] &\leq n^{-2} h_{n+1}^{-m} \int \left| \sum_{j=1}^n [e^{iu'X_j} - \phi(u)] \frac{k(h_j u)}{k(h_n u)} \right|^2 |k(h_n u)| du \\ &\quad + n^{-2} h_{n+1}^{-2m} \int |k(u)| du. \end{aligned}$$

Adding and subtracting  $Y_n$  we obtain

$$\begin{aligned} E[Y_{n+1} | F_n] &\leq Y_n - n^{-2} (h_n^{-m} - h_{n+1}^{-m}) \int \left| \sum_{j=1}^n [e^{iu'X_j} - \phi(u)] \frac{k(h_j u)}{k(h_n u)} \right|^2 |k(h_n u)| du \\ &\quad + n^{-2} h^{-2m} \int |k(u)| du. \end{aligned}$$

Let,

$$(15) \quad V_n = (nh_{n+1}^{2m})^{-1} \int |k(u)| du$$

and

$$(16) \quad U_n = n^{-2}(h_n^{-m} - h_{n+1}^{-m}) \int \left| \sum_{j=1}^n [e^{iu'X_j} - \phi(u)] \frac{k(h_j u)}{k(h_n u)} \right|^2 |k(h_n u)| du.$$

Note that  $V_n$  is not a random variable. Hence  $EV_n = V_n$ , and, using condition (10), we have  $\sum_{n=1}^{\infty} E|V_n| = \sum_{n=1}^{\infty} |V_n| < \infty$ . Thus, in order to use Lemma 1 we need to show now that  $\sum_{n=1}^{\infty} E|U_n| < \infty$ . From (16) we have, taking expectations,

$$(17) \quad \begin{aligned} E|U_n| &= n^{-2} |h_n^{-m} - h_{n+1}^{-m}| \int (1 - |\phi(u)|^2) \sum_{j=1}^n \left| \frac{k(h_j u)}{k(h_n u)} \right|^2 |k(h_n u)| du \\ &\leq n^{-2} |h_n^{-m} - h_{n+1}^{-m}| h_n^{-m} \int n |k(u)| du. \end{aligned}$$

By using the fact that  $(1-t^m) = \left(\sum_{j=1}^m t^{j-1}\right)(1-t)$  and  $h_n h_{n+1}^{-1} \rightarrow 1$ , Van Ryzin (1969) shows

$$|h_{n+1}^{-m} - h_n^{-m}| \sim m h_n^{m-1} |h_{n+1}^{-1} - h_n^{-1}|.$$

Thus the upper bound in (17) is asymptotically equivalent to  $\frac{m}{n h_n^{2m-1}} |h_{n+1}^{-1} - h_n^{-1}| \times \int |k(u)| du$ . Hence, using condition (11),  $\sum_{n=1}^{\infty} E|U_n| < \infty$ . Thus conditions (iii) and (iv) of Lemma 1 are satisfied, where  $Y'_n = U_n + V_n$ .

The conditions of Lemma 1 are satisfied for  $Y_n$  and so  $Y_n$  tends to zero almost everywhere. Thus,

$$\sup_x |\hat{f}_{n+1}(x) - E\hat{f}_{n+1}(x)| \rightarrow 0 \quad \text{a.s.} \quad \text{as } n \rightarrow \infty.$$

Now,

$$\begin{aligned} \sup_x |E\hat{f}_n(x) - f(x)| &= \sup_x \left| \int \left( n^{-2} \sum_{j=1}^n K_j(u) f(x-u) - n^{-1} \sum_{j=1}^n K_j(u) f(x) \right) du \right| \\ &\leq \sup_x \int n^{-1} \left| \sum_{j=1}^n K_j(u) \right| |f(x-u) - f(x)| du \\ &\leq \sup_x \int_{\|u\| \leq \delta} n^{-1} \left| \sum_{j=1}^n K_j(u) \right| |f(x-u) - f(x)| du \\ &\quad + \sup_x \int_{\|u\| > \delta} |f(x-u) - f(x)| n^{-1} \sum_{j=1}^n K_j(u) du \\ &\leq \sup_x \sup_{\|u\| \leq \delta} |f(x-u) - f(x)| + 2 \sup_x n^{-1} f(x) \sum_{j=1}^n \int_{\|u\| > \delta/h_j} K(u) du \end{aligned}$$

Since  $f(x)$  is uniformly continuous, the first term can be made arbitrarily small by choosing  $\delta$  sufficiently small. In the second term, notice that for this fixed  $\delta$

$$f(x) \int_{\|u\| \geq \delta/h_n} K(u) du \rightarrow 0$$

as  $n \rightarrow \infty$  at all points  $x$ . Hence the Cesàro sum approaches zero and the proof is complete.

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